



## Abstract

Although the non-uniqueness of the solution is commonly mentioned in the context of studies that perform spectral decompositions to separate source and propagation effects, its impact on the interpretation of the results is often overlooked. The purpose of this study is to raise awareness on this important subject for modelers and users of the models and to evaluate the impact of strategies commonly applied to constrain the solution. In the first part, we study the connection between the source-station geometry of an actual data set and the properties of the design matrix that defines the spectral decomposition. We exemplify the analyses by considering a geometry extracted from the data set prepared for the benchmark Community Stress Drop Validation Study (Baltay et al., 2021). In the second part, we analyze two different strategies followed to constrain the solutions. The first strategy assumes a reference site condition where the average site amplification for a set of stations is constrained to values fixed a-priori. The second strategy consists in correcting the decomposed source spectra for unresolved global propagation effects. Using numerical analysis, we evaluate the impact on source scaling relationships of constraining the corner frequency of magnitude 2 events to 30 Hz when the true scaling deviates from this assumption. We show that the assumption can not only shift the overall seismic moment versus corner frequency scaling but can also affect the source parameters of larger events and modify their spectral shape.

## Plain Language Summary

Source properties at seismogenic depths cannot be measured directly. Therefore, the characterization of the size and strength of the source has to rely on our ability to isolate the footprint left by the source into seismic recordings acquired at the Earth's surface. Since propagation effects are not known in detail, questions about absolute measure of the earthquake source have not unique answers. An approach in which source, propagation and site effects are isolated within observations is called a decomposition approach. In this study, we discuss the non-uniqueness of the decomposition and the impact of the trade-offs affecting the different decomposed terms. Using a data set collected in Southern California during the 2019 Ridgecrest sequence, we connect non-uniqueness and trade-offs to properties of matrices describing the problem such as the design, covariance and resolution matrices. We also develop numerical analysis to quantify the impact on source scaling of the constraints used to resolve the non-uniqueness. Overall, this study raises awareness on the consequences that different strategies applied to deal with the non-uniqueness can have on the final outcomes of the decomposition, an impact that should not be overlooked when comparing conclusions drawn from different studies performed in the same area.

## 1 Introduction

The retrieval of the moment rate function or spectrum of an earthquake from a set of recordings requires the solution of an inverse problem. In the model framework where the seismic waves propagate through the Earth's interior as elastic waves, the transfer function of the system describes the effects of seismic wave propagation from the hypocenter to the stations, including near-surface site effects. Therefore, under the assumption of linear and time-invariant systems, the source function retrieval is a deconvolution problem (Helmberger & Wiggins, 1971; Hartzell, 1978; Mueller, 1985; Bertero et al., 1997). Two main limitations affect the retrieval process: first, propagation effects are also unknown, introducing a system identification problem; second, the unknown propagation effects filter the input signal and source-related information can be recorded only within a limited bandwidth. The consequence of the first issue is that solving the inverse problem involves either the determination of both the source and the propagation terms (Andrews, 1986) or requires assumptions about the propagation effects, such as within the empir-

ical Green’s function approach (e.g., Hartzell, 1978; Bertero et al., 1998, among others). Regarding the filtering issue, since source characteristics outside the data bandwidth cannot be reconstructed only using observations (Ide & Beroza, 2001), the detection of high-frequency source parameters for small events is hampered by the limited bandwidth and by the competing source and attenuation effects. The detection limits of the source parameters for small events have been investigated in a number of studies using numerical simulations and theoretical models (e.g., Kwiatek & Ben-Zion, 2016; Chen & Abercrombie, 2020; Bindi, Spallarossa, et al., 2020). Although several factors contribute to determine the magnitude threshold (e.g., source-station geometry, severity of the near surface attenuation, stress drop values for small events, sampling rate, among others), the determination of reliable source parameters below magnitude 2 is challenging with surface monitoring networks.

The analysis of data sets with a high level of redundancy in the cross-sampling between stations and events has been proposed to simultaneously determine source, propagation and site effects in the spectral domain (Andrews, 1986; Castro et al., 1990; Boatwright et al., 1991; Shearer et al., 2006; Oth et al., 2011; Bindi, Zaccarelli, & Kotha, 2020; Shible et al., 2022). The possibility to observe the same source spectrum at different distances and azimuths, with each site term combined with different sources, allows to set up an overdetermined system of equations that can be solved in a least-squares sense to isolate the different terms. A peculiarity of such an approach is that the solution is not unique since, regardless of the number of available recordings, the model assumptions, and the wavefield selected for analysis (e.g., S-waves or coda), the design matrix is rank deficient (Andrews, 1986). Specific solutions are obtained by imposing a-priori constraints to the solution, causing difficulties in comparing results from different studies even when performed in the same area. In addition to the non-uniqueness generated by the convolution model, other factors limit our ability to resolve source, propagation and site effects from recordings. We identify two major limitations: the first limitation is connected to the physics of the seismological problem, where the different terms show some degree of correlation (for example, the same high-frequency spectral content could be obtained by different combinations of source parameters and near-surface attenuation); the second limitation depends on the assumptions made over the different terms. Regarding the latter, examples are the assumption of dealing with spherically averaged source spectra (i.e., neglecting radiation pattern and directivity effects) and of assuming a 1D spectral attenuation depending only on hypocentral distance (i.e., neglecting lateral velocity and attenuation variabilities; not accounting explicitly for depth dependencies), which are typical assumptions for decomposition studies. The impact of these two hampering factors can be exacerbated or mitigated by the geometry of the analyzed problem (translated into characteristics of the design matrix) and, in general, it depends on frequency. Moreover, the impact depends on the analyzed wavefield since, for example, seismic coda is less sensitive to azimuthal variability of the source radiation than direct waves (Aki & Chouet, 1975; Aki, 1981; Mayeda & Malagnini, 2010).

The target of this study is to discuss the connections between the source-station geometry of an actual data set and the properties of the design matrix defining the source, propagation and site spectral decomposition. In particular, this study is motivated by the ongoing effort to develop a community stress drop initiative for comparing the source parameters estimated for the shared Ridgecrest data set (Baltay et al., 2021; Trugman, 2020) by applying a wide variety of approaches. Refreshing the connection between the non-uniqueness of the spectral decomposition approach and the trade-offs affecting the source, propagation and site terms can support the comparison of spectral decomposition results with other approaches, and increase awareness of the impact of non-uniqueness of the solution on the interpretation of results (see also Chen & Abercrombie, 2020; Zhang et al., 2022). For this reason, we also discuss strategies applied to select specific source spectra from the infinite set of possible solutions, evaluating their impact on source parameters estimation.

## 2 Spectral decomposition of the Fourier amplitude spectra

The Generalized Inversion Technique (GIT) aims at isolating the source, propagation and site terms from the Fourier Amplitude Spectra (FAS) of the recorded P- or S-waves. Large data sets with a high level of redundancy in the station and earthquake sampling are decomposed under the assumption that source, propagation and site effects combine through a convolution product expressed in the Fourier domain as

$$O_{ij}(f) = S_i(f) + P(R_{ij}, f) + Z_j(f) \quad (1)$$

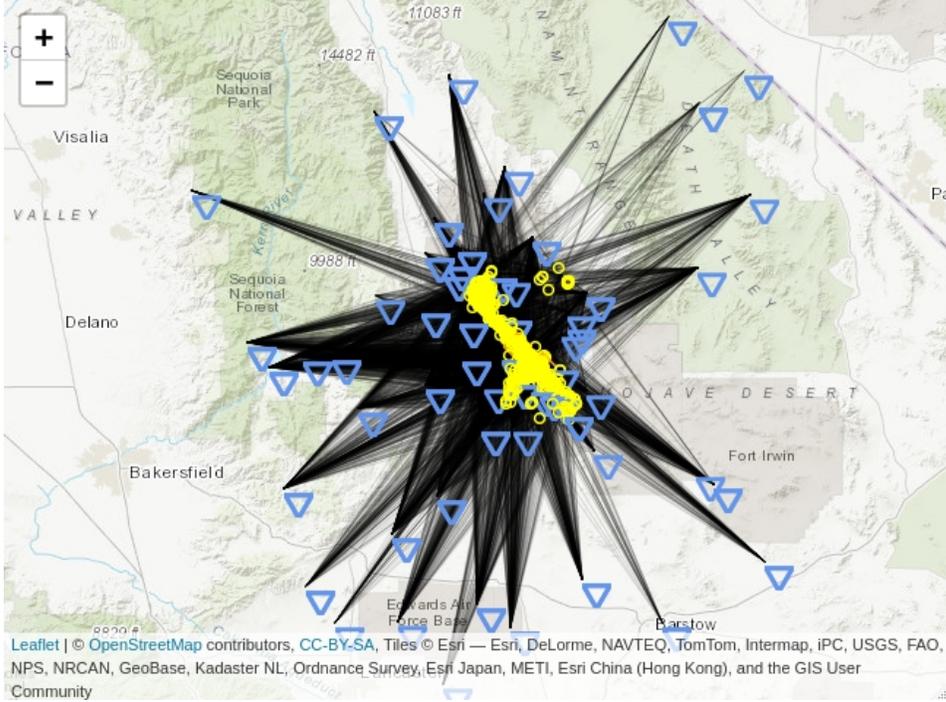
where the  $O_{ij}(f)$  is the logarithm of the spectral amplitude at frequency  $f$  for earthquake  $i$  recorded at station  $j$ ;  $S_i(f)$  is the logarithm of the source spectrum for event  $i$ ;  $P(R_{ij}, f)$  is the logarithm of the propagation term for the hypocentral distance  $R_{ij}$ ;  $Z_j(f)$  is the logarithm of the site term for station  $j$ . In equation 1, the propagation effects are assumed to be isotropic and controlled by the hypocentral distance. Two strategies can be applied to isolate source, propagation and site effects: the parametric (Castro et al., 1990; Boatwright et al., 1991) and the non-parametric (Andrews, 1986) approaches. In the former, source and attenuation terms are described in terms of standard seismological models, such as a parametrization of the source spectra in terms of seismic moments and corner frequencies and a parametrization of the propagation in terms of geometrical spreading and  $t^*$ ; in the non-parametric approach, equation 1 is solved without imposing a-priori parametric models on the different terms. In this study, we focus only on the non-parametric approach, although the non-uniqueness of the solution is intrinsic to both the parametric and non-parametric approaches and, more in general, to any decomposition approach applied in the time or frequency domains to isolate source, propagation and site contributions to recorded time histories or Fourier spectra. By analyzing each frequency separately, equation 1 generates a set of overdetermined linear systems that can be solved in a least-squares sense. In the following, we provide a description of the coefficient matrix considering an extraction of the data set based on the 2019 Ridgecrest sequence (Trugman, 2020) compiled for the Community Stress Drop Validation study (Baltay et al., 2021). We consider 554 earthquakes recorded by 94 stations (counting co-located sensors separately) generating 11064 spectral amplitudes at  $f=4.2$  Hz; magnitudes are in the range 2.5-7.1 and hypocentral distances cover the range 5-111.4 km. For the attenuation term  $P(R, f)$ , the range of distances from 4 to 112 km is discretized by considering 37 nodes spaced 3 km apart. The geometry of the data set is shown in Figure 1.

## 3 Matrix of coefficients, covariance and resolution

The design matrix of the linear system generated by equation 1 is a sparse matrix with in general four non-zero elements for each row. The matrix has dimension  $M \times N$ , where  $M$  is the number of recordings and  $N$  is the sum of the number of events  $N_e$ , of stations  $N_s$  and the number of nodes  $N_d$  used to discretize the distance range. The structure of the matrix can be expressed in terms of dummy variables used to select the event, station and distance bin relative to each recording. If the recording in the  $m$ -th row is associated with the event  $i$  recorded by station  $j$  and the distance  $R_{ij}$  lies between nodes  $k$  and  $k+1$  (i.e.,  $R_k \leq R_{ij} < R_{k+1}$ ), equation 1 can be written as

$$O_{m,n} = \delta_{n,i} S_i + \delta_{n,N_e+j} Z_j + (1 - w_k) \delta_{n,k+N_e+N_s} P_k + w_k \delta_{n,k+1+N_e+N_s} P_{k+1} \quad (2)$$

where  $S_i$  with  $i = 1, \dots, N_e$ ,  $Z_j$  with  $j = 1, \dots, N_s$ , and  $P_k$  with  $k = 1, \dots, N_d$  define the vector of unknowns. In our example, since  $N_e=554$ ,  $N_s=94$ , and  $N_d=37$ , the total number of unknowns for a given frequency is  $N=685$ , and the number of data counted at  $f=4.2$  Hz is  $M=11064$ . In equation 2, the attenuation term is linearly interpolated between nodes  $k$  and  $k+1$ , i.e.,  $w_k = (R_{ij} - R_k)/(R_{k+1} - R_k)$ . Therefore, all entries in each row of the design matrix are zero except those for one specific column in the event-column group, one specific column in the station-column group and two adjacent columns

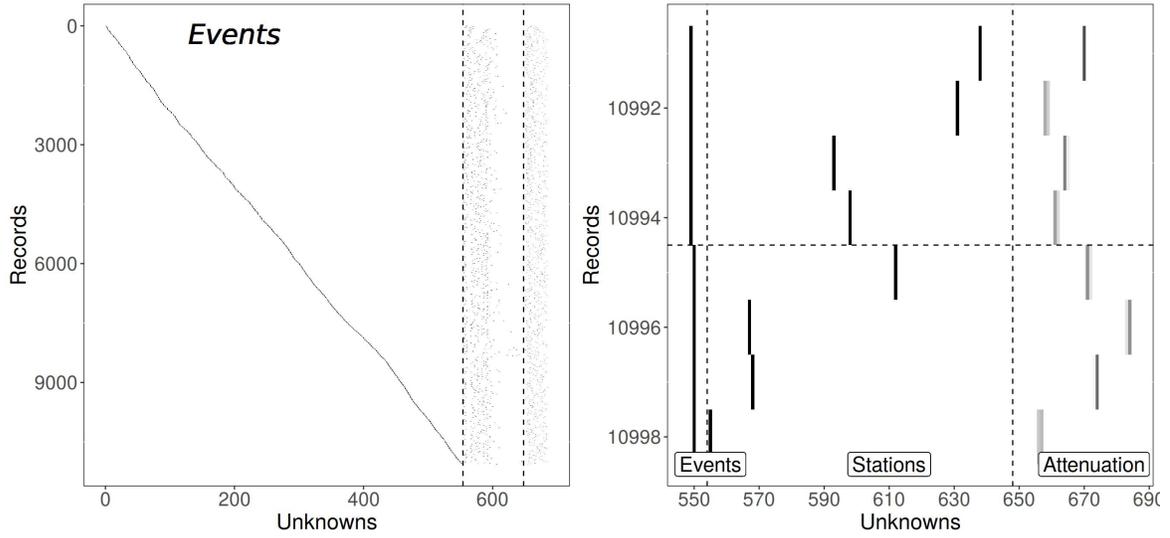


**Figure 1.** Event-station geometry of the data set considered in this study, extracted from the Community Stress Drop Validation study (Baltay et al., 2021): black lines connect epicenters (circles) of the considered earthquakes with the location of the recording stations (triangles)

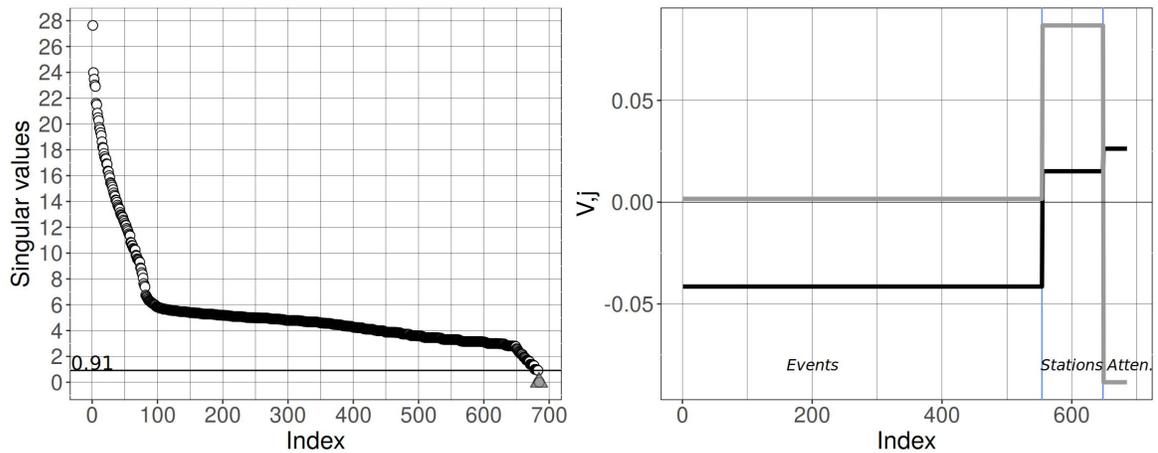
171 in the attenuation-column group (the latter reduce to one if the distance  $R_{ij}$  is coinci-  
 172 dent with one of the nodes  $R_k$ ). The design matrix for the geometry in Figure 1 is shown  
 173 in Figure 2, where the recordings are ordered per event (all recordings of the same event  
 174 are consecutive). This matrix is based on the data available at 4.2 Hz. In the lower panel  
 175 of Figure 2, which shows a detail of the matrix for rows 10991 to 10998 corresponding  
 176 to two different earthquakes, non-null elements are shown by vertical bars (the value is  
 177 1 for the event and station columns, two values between 0 and 1 and summing to 1 for  
 178 the two adjacent attenuation columns). The least-square solution of system 1 is not unique  
 179 because there are two unresolved degrees of freedom: since we are summing three terms,  
 180 we can add a constant to one term and remove the same constant from another and the  
 181 sum will not change (Andrews, 1986). To test the rank deficiency of the design matrix,  
 182 we perform its singular value decomposition

$$G = USV^T \quad (3)$$

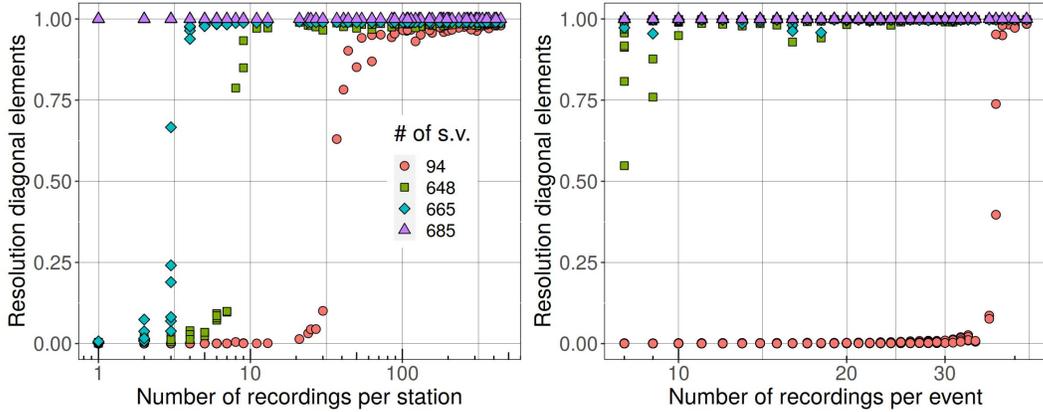
183 where  $G$  is the  $M \times N$  design matrix,  $U$  is the  $M \times M$  orthogonal matrix whose columns  
 184 generate the data space;  $V$  is the  $N \times N$  orthogonal matrix whose columns generate the  
 185 model space;  $S$  is an  $M \times N$  diagonal matrix with non-negative diagonal elements called  
 186 singular values. The singular values for the geometry of Figure 1 are shown in Figure  
 187 3. As expected, among the 685 singular values, two are numerically close to zero (i.e. of  
 188 the order of  $10^{-16}$ ), confirming that the model null space (kernel) has dimension (nul-  
 189 lity) equal to 2. The right panel of Figure 3 shows the columns  $V_{684}$  and  $V_{685}$  of the ma-  
 190 trix  $V$ , which are associated with the null singular values (i.e., the singular values are  
 191 ordered in decrescent order). These two columns form an orthonormal basis for the ker-  
 192 nel of the design matrix. Each element of the basis consists of a constant value on each  
 193 column block related to event, station and attenuation, with the sum of the three con-  
 194 stant values equal to 0. They represent the trade-off existing among the source, station



**Figure 2.** Design matrix (left) for the network geometry shown in Figure 1. Columns from 1 to 554 are relevant to events; columns from 555 to 648 to stations; columns from 649 to 685 to attenuation. The right panel shows a detail of the design matrix for selected rows and columns.



**Figure 3.** Left. Singular values of the design matrix compiled for the data set in Figure 1; the two null singular values are shown as gray triangles. Right. Base vectors for the null space corresponding to columns  $V_{684}$  and  $V_{685}$ , where  $V$  is the right orthonormal matrix of the SVD decomposition (equation 3).

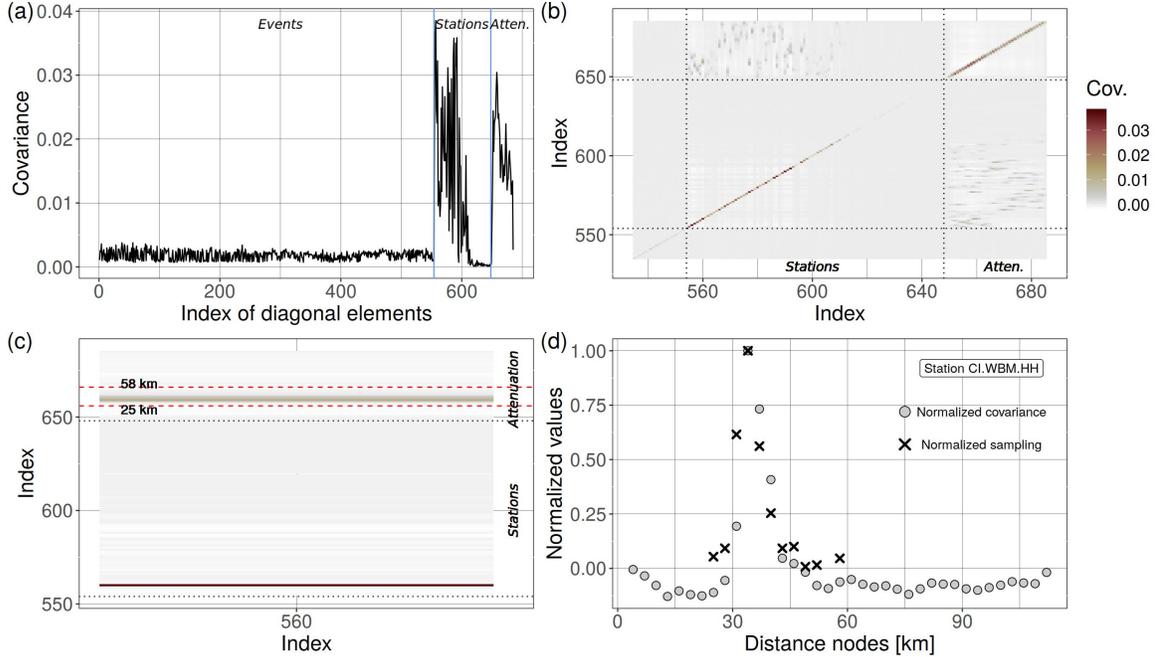


**Figure 4.** Diagonal elements of the resolution matrix versus number of recordings for each station (left) and event (right) considering different cutoff thresholds applied to singular values.

195 and propagation terms and without additional information (constraints), the solution  
 196 of the system 1 is not unique. This point is further discussed later. The non-null singular  
 197 values vary from 0.91 to 27.63 (Figure 3). To investigate the connection between singular  
 198 values and the resolution of the model parameters, Figure 4 shows the diagonal  
 199 elements of the resolution matrix  $R$  computed considering only the largest  $q$  singular val-  
 200 ues (truncated SVD):

$$R = V_q V_q^T \quad (4)$$

201 where  $V_q$  is the  $N \times q$  matrix composed by the first  $q$  columns of  $V$  (equation 3). For  $q=n=685$ ,  
 202  $R$  is the identity matrix (overdetermined system), as well as for  $q=684$  (since the last  
 203 two singular values are zero). For  $q < 684$ , the truncation removes model space basis vec-  
 204 tors associated with small singular values (regularization). In Figure 4, we consider the  
 205 cases  $q=94$ , 648, and 665 where the first two cut-offs correspond approximately to the  
 206 sharp changes in the convexity of the distribution of singular values (Figure 3, left). Fig-  
 207 ure 4 shows that the regularization affects mostly those model parameters (station on  
 208 the left and events on the right) with low sampling; the first 94 elements allow to resolve  
 209 well model parameters with more than 40 recordings; when the number of singular val-  
 210 ues is increased to 648, parameters with at least about 10 recordings are reasonably well  
 211 resolved; considering 665 singular values, model parameters with at least 5 recordings  
 212 are resolved. In solving system 1 a threshold is generally applied on the minimum num-  
 213 ber of recordings per station and per event. For example, for the decomposition performed  
 214 in the framework of the Stress Drop benchmark, we set the threshold to 6. Other im-  
 215 portant information associated with the GIT design matrix is provided by the covari-  
 216 ance matrix (Boatwright et al., 1991; Bindi et al., 2006). Figure 5(a) shows the diago-  
 217 nal elements of the covariance matrix, i.e., the sample variances of the model paramet-  
 218 ers. Since the elements of the design matrix in the event and station columns are ei-  
 219 ther 0 or 1, and since the number of rows  $M$  is much larger than the total number of ones  
 220 in each column, the variances increase almost linearly with the sampling of each station  
 221 or event. The off-diagonal elements of the covariance matrix provide information about  
 222 the trade-offs (correlation) between different model parameters that jointly contribute  
 223 to generating the observations. These trade-offs are determined by the geometry of the  
 224 problem. Figure 5(b) shows that there is a trade-off between station and attenuation model  
 225 parameters. For example, Figure 5(c) focuses on model parameter 560, corresponding  
 226 to station CI.WBM.HH. The recordings available for this station are in the distance range  
 227 of 25 to 50 km, with median and mode hypocentral distance equal to about 35 km. Fig-  
 228 ure 5(d) confirms that there is a trade-off between the station term and the attenuation  
 229 coefficients in the distance range sampled by the station, in particular for the mostly sam-



**Figure 5.** (a) Diagonal elements of the covariance matrix. (b) detail of the covariance matrix considering the station and attenuation columns. (c) detail of the covariance matrix for station CI.WBM.HH (column 560) considering only the station and attenuation portion of the matrix. Red dashed lines indicate the distance range sampled by recordings at CI.WBM.HH. (d) circles indicate the normalized covariance off-diagonal entries between station CI.WBM.HH and the attenuation columns; crosses indicate the normalized number of recordings for station CI.WBM.HH available within each distance bin.

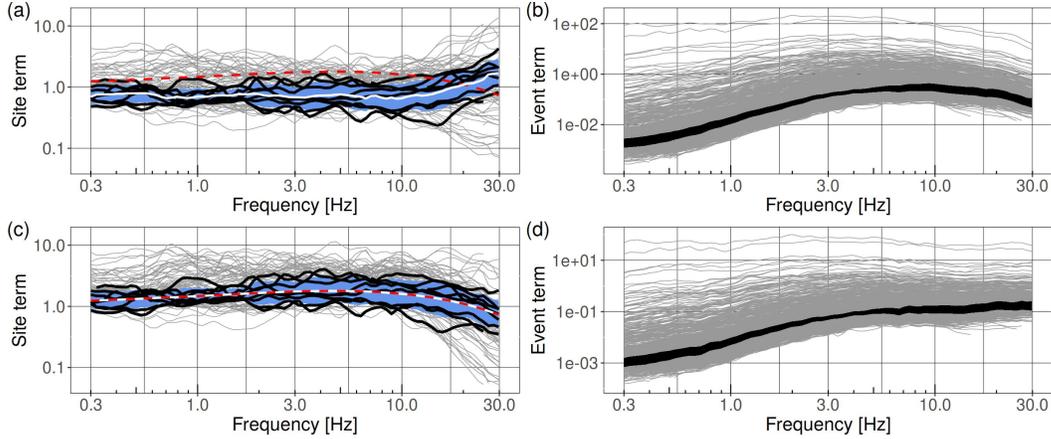
230 plied distances. In order to limit the trade-offs, it is important that different stations sam-  
 231 ple the same distance bins, avoiding that a few stations dominate specific distance in-  
 232 tervals, and that each station provides recordings over wide distance and magnitude in-  
 233 tervals (Shible et al., 2022).

#### 234 4 Approaches to remove the non-uniqueness of the solution

235 Different strategies can be followed to select a specific solution among the infinite  
 236 possible ones. In the following, we discuss two widely used approaches: the first one con-  
 237 sists in constraining the site term within the design matrix; the second one makes as-  
 238 sumptions on the source parameters for some reference events that in turn are used to  
 239 correct the source spectra provided by the unconstrained decomposition.

##### 240 4.1 Constraining site and attenuation terms

241 The first strategy consists in adding some rows to the design matrix, forcing the  
 242 solution to satisfy specific conditions that make the matrix full rank. A typical choice  
 243 is to constrain to zero the average of the logarithm of all site amplifications. In this way,  
 244 the site effects do not bias the observations on average but contribute to the overall vari-  
 245 ability. A similar constraint is often applied when calibrating a local magnitude scale us-  
 246 ing data a seismological network (with stations mostly installed on rock) by setting to  
 247 zero the average of the station magnitude corrections (Bakun & Joyner, 1984; Savage



**Figure 6.** Impact on the source spectra of constraints applied to the site term. a) Site amplifications obtained by setting to 1 the average of all site terms. b) source spectra relevant to the site terms as in panel a. c) site amplifications obtained by constraining the average (white curve) of selected rock sites (black curves) to be identical to the spectral function shown as a dashed line; the same reference sites (black) and their average amplification (white)  $\pm$  one standard deviation are also shown in panel a. d) source spectra relevant to the site terms in panel c. In panels b and d, the black curves are source spectra of magnitude 3 events.

248 & Anderson, 1995; Langston et al., 1998; Baumbach et al., 2003). Since there are two  
 249 null singular values, a second constraint is applied to the attenuation by requiring that  
 250 the attenuation assumes a given value at a reference distance, i.e.  $\log P_k=0$  for  $R_k = R_{ref}$ .  
 251 As a consequence, the source spectra is scaled at  $R_{ref}$ . For the case study analyzed in  
 252 this study, we compare the impact of applying two different constraints on the site term,  
 253 and we set to 0 the logarithm of the attenuation at 10 km. The first site constraint co-  
 254 incides with the standard requirement that the average of the logarithm site amplifica-  
 255 tions is 0 for all frequencies. As expected, this requirement together with the constraint  
 256 applied to the attenuation term, removes the two null singular values and the condition  
 257 number of the design matrix is now  $\sim 2543$ . Since the considered stations are installed  
 258 in different geological settings, the site amplifications (Figure 6a) show a large variabil-  
 259 ity, in particular above 10 Hz. The applied constraint implies that the site terms rep-  
 260 resent the site amplifications with respect to the network average. If the average ampli-  
 261 fication deviates from the imposed flat spectral behavior, the deviation from the true aver-  
 262 age is moved to the source spectra (Figure 6b). The acceleration source spectra show  
 263 an average high frequency decay different from the flat asymptote predicted by the omega-  
 264 square source model, and the absorbed average site amplification contributes to the ob-  
 265 served high-frequency spectral fall-off. In Figure 6a, the amplification of 6 stations in-  
 266 stalled on rock (vs30 values measured or inferred from geology above 710 m/s) are shown  
 267 in black. They have a flat amplification lower than the network average at frequencies  
 268 smaller than about 10 Hz, and then they show a positive amplification trend with fre-  
 269 quency. The positive trend indicates that the near-surface attenuation for these stations  
 270 ( $k_0$  parameter, (Anderson & Hough, 1984)) is probably weaker than the average near  
 271 surface attenuation of the overall network. Therefore, following Bindi, Zaccarelli, and  
 272 Kotha (2020), in the second reference site strategy we constrain the average amplifica-  
 273 tion of the six stations to the crustal amplification proposed by Campbell and Boore (2016)  
 274 for the National Earthquake Hazards Reduction Program (NEHRP) B/C boundary mul-  
 275 tiplied by an exponential term with  $k_0=0.034$  s (Figure 6c, dashed line). The source spec-  
 276 tra obtained using this constraint (Figure 6d) show an average flat high frequency spec-

277 tral level, although single spectra have still small positive and negative slopes ( $k_{source}$ )  
 278 that have been shown to correlate well with the between-event ground motion variabil-  
 279 ity at high frequencies (Bindi, Zaccarelli, & Kotha, 2020). If we compare the logarithm  
 280 of the ratio between the average spectra for magnitude 3 events with the logarithm of  
 281 the inverse average site amplifications for the reference stations, they are identical (Fig-  
 282 ure S1 of the Supplements). This confirms that the constraint applied to the site am-  
 283 plifications breaks the trade-off between source and site terms but the constrained so-  
 284 lutions provide the same predictions as the unconstrained solutions and the analysis of  
 285 the residuals alone is not sufficient to discriminate among the quality of solutions ob-  
 286 tained by applying different constraints.

## 287 4.2 Reference source conditions

288 A different approach followed in the context of source studies is to correct the ob-  
 289 tained source spectra  $S_j(f)$  after the decomposition. The correction is performed by us-  
 290 ing additional information about source parameters and making assumptions about the  
 291 source spectral shape (e.g., Baltay et al., 2011; Walter et al., 2017; Trugman & Shearer,  
 292 2017; Morasca et al., 2022). For example, if the source spectra are assumed to be omega-  
 293 square, the source spectra can be scaled by fixing the seismic moment and the corner  
 294 frequency of one or more calibration events (reference source condition). In the follow-  
 295 ing, we discuss some aspects connected to recently applied reference source conditions  
 296 (e.g., Shearer et al., 2022). It is worth noting that if the source spectra are described in  
 297 terms of a spectral model, then the model can be introduced in equation (1) making the  
 298 GIT approach parametric. In this case, the system can be solved to directly determine  
 299 the source parameters (e.g. seismic moment and corner frequency) and the constraint  
 300 on seismic moment and corner frequency can be included directly in the design matrix  
 301 (e.g., Moya & Irikura, 2003). In this case, the complexity of the source models gener-  
 302 ally makes the system non linear.

## 303 5 Impact of constraining the corner frequency of calibration events

304 In the reference source strategy, reference moment magnitudes are used to remove  
 305 the bias on the estimated seismic moments (shift of the source spectra), whereas an Em-  
 306 pirical Correction Spectrum (ECS) is determined to correct the shape of the decomposed  
 307 source spectra (e.g., Shearer et al., 2006; Trugman & Shearer, 2017; Shearer et al., 2022).  
 308 The spectral shape of the source spectra is described in terms of a standard model (e.g.,  
 309  $\omega$ -square) and the corner frequency of selected events is assumed to be known. Typically,  
 310 the corner frequency is fixed for small events, hereinafter referred to as global empiri-  
 311 cal Green’s functions (gEGF). For example, in recent applications for southern Califor-  
 312 nia, the corner frequency of the gEGF with local magnitude 1.5 was fixed to 30 Hz (Shearer  
 313 et al., 2022) whereas Chen and Abercrombie (2020); Zhang et al. (2022) let the data to  
 314 select the constant value.

315 We consider a simplified approach where the errors arising from the decomposition  
 316 are neglected, as well as the propagation of experimental errors from data to solutions.  
 317 We assume that the source spectra isolated by the spectral decomposition are given by  
 318 the product between the true source spectra  $S$  and the unknown ECS. The latter is es-  
 319 timated by considering small earthquakes (also called calibration events) representing  
 320 the gEGF, whose corner frequency is a-priori fixed to a certain value, that is:

$$321 ECS_{est} = \frac{gEGF \cdot ECS}{g\widehat{EGF}(f_c = \hat{f}_c; shape = \omega^2)} \quad (5)$$

322 In equation 5, the term  $gEGF \cdot ECS$  represents the  $gEGF$  source spectrum as provided  
 323 by the decomposition; therefore, an estimate of the ECS (indicated as  $ECS_{est}$ ) is ob-  
 324 tained by removing  $g\widehat{EGF}$  from this term. The quantity  $g\widehat{EGF}$  represents the a-priori  
 assumed spectrum for the  $gEGF$  (in our case, an omega-square spectrum with corner

325 frequency fixed to  $\hat{f}_c$  and a known seismic moment). The  $ECS_{est}$  spectrum is then used  
 326 to correct the spectra of the target events provided by the decomposition and get their  
 327 source spectra, that is:

$$S_{est} = \frac{S \cdot ECS}{ECS_{est}} \quad (6)$$

328 where  $S$  is the true source spectrum of the target event;  $S \cdot ECS$  is the source spectrum  
 329 of the target event as provided by the decomposition;  $S_{est}$  is the estimated source spec-  
 330 trum of the target event. Substituting equation 5 into equation 6, we obtain

$$S_{est} = \frac{g\widehat{EGF}_{\hat{f}}}{gEGF} S \quad (7)$$

331 where  $g\widehat{EGF}_{\hat{f}}$  is the spectrum assumed for the calibration event and  $gEGF$  its true spec-  
 332 trum. Equation 7 shows that in our numerical tests the shape of the ECS does not play  
 333 any role (since it affects the target and the calibration spectra in the same way) and the  
 334 quality of the retrieved target source spectrum is controlled by the ratio between the as-  
 335 sumed and the true spectra of the calibration event. Even if the shape of the ECS is not  
 336 entering in equation 7, in order to show its impact on the ECS and source target esti-  
 337 mations (as in equations 5 and 6) in the following we assume an exponential shape for  
 338 describing the ECS:

$$ECS(f) = e^{\pi k f} \quad (8)$$

339 The spectral shape of equation 8 resembles typical shapes of ECS shown in literature (e.g.,  
 340 Shearer et al., 2006, 2022).

## 341 5.1 Numerical Test

342 To evaluate the impact of assuming  $f_c = \hat{f}_c$  for the gEGF, we perform a numer-  
 343 ical test considering: 1) a synthetic distribution of events; 2) uncertainty on seismic mo-  
 344 ment; 3) variability of the stress drop for the same seismic moment. The components  
 345 are the following:

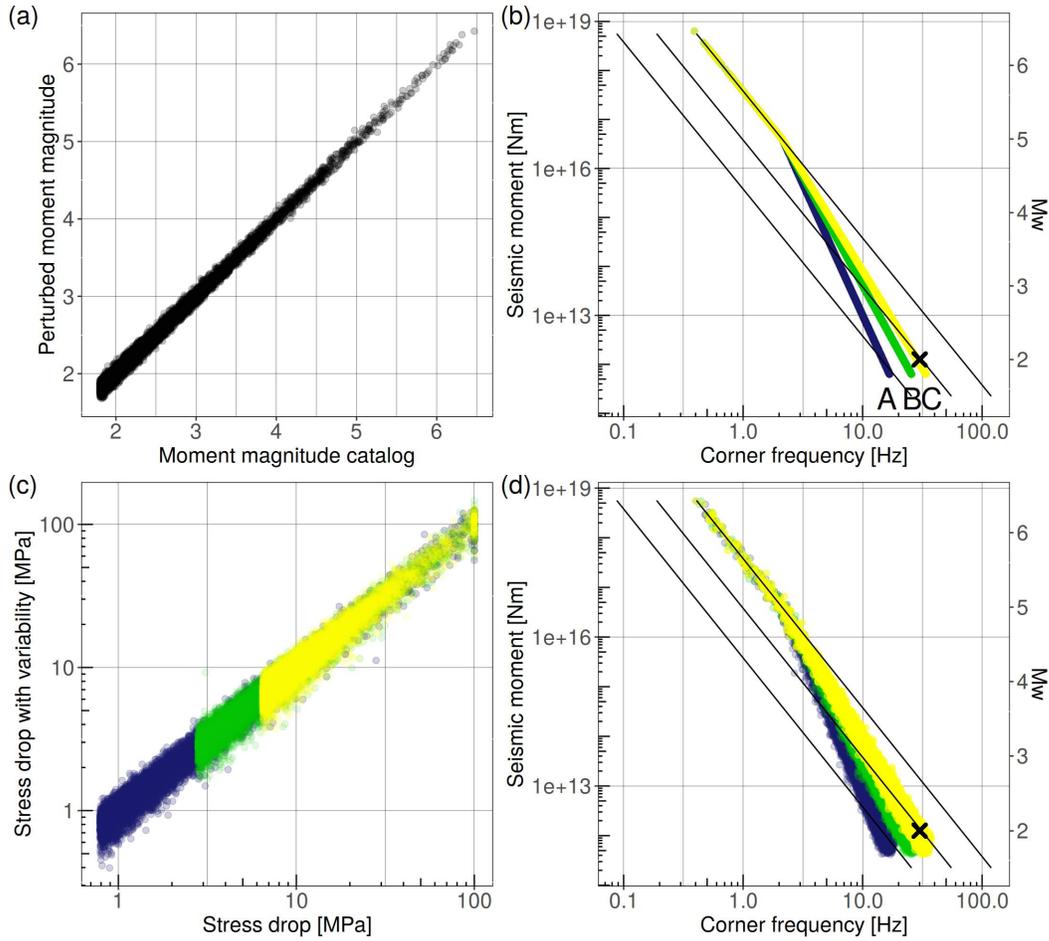
- 346 • Event distribution. We generate a distribution of events following a Gutenberg-  
 347 Richter magnitude distribution with minimum magnitude  $M_{min}=1.8$  and max-  
 348 imum magnitude  $M_{max}=6.5$ ; we generate a catalog of 10000 earthquakes with b-  
 349 value equal 0.6 (we set the b to a low value to increase the number of larger events).  
 350 The distribution is shown in Figure S2 of the Supplements.
- 351 • Uncertainty on magnitude. We perturb the magnitude values by adding random  
 352 numbers extracted from a normal distribution centered on the true magnitude value  
 353 and with standard deviation equal to 0.05 magnitude units.
- 354 • Variability on  $\Delta\sigma$ . We allow the stress drop to vary for the same seismic moment.  
 355 The applied perturbation is as follows
  - 356 1. for each  $Mw$ , we compute  $Mw_{err} = \mathcal{N}(Mw, 0.05)$  (Figure 7a); from  $Mw$  and  
 357  $Mw_{err}$ , we compute the seismic moments  $Mo$  and  $Mo_{err}$ , respectively.
  - 358 2. we assume three different scaling models, referred to A, B and C, between seis-  
 359 mic moment and corner frequency as shown in Figure 7b; each scaling corre-  
 360 sponds to a different dependency of the stress drop  $\Delta\sigma$  on seismic moment, be-  
 361 ing scaling A deviating with the largest error and scaling C being compatible  
 362 on average with the gEGF assumption.
  - 363 3. for each scaling, we compute the true corner frequency from  $Mw$  and  $\Delta\sigma$  us-  
 364 ing the formula for circular rupture model with uniform stress drop (Brune, 1970,  
 365 1971);
  - 366 4. we add a variability over  $\Delta\sigma$  as  $\Delta\sigma_{var} = \mathcal{N}(\Delta\sigma, 0.15*\Delta\sigma)$  (Figure 7c) . The  
 367 normal distribution is truncated below  $0.5*min(\Delta\sigma)$  and above  $2*max(\Delta\sigma)$ .

368 Figure 7d shows the scaling considering the errors on  $M_o$  and the variability on  
 369  $\Delta\sigma$ .

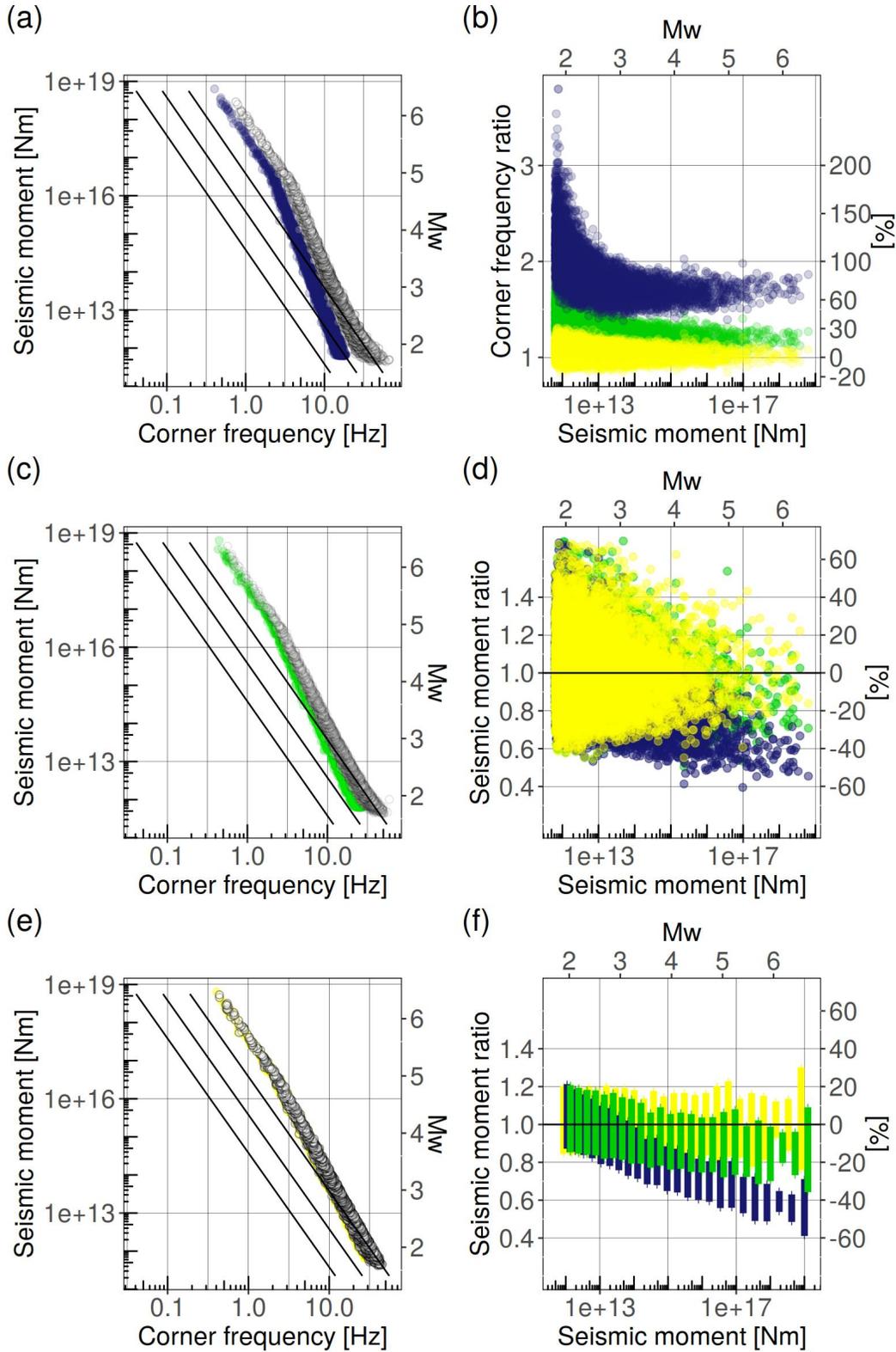
- 370 • average EGF. Following previous studies (e.g. Shearer et al., 2022), a set of candi-  
 371 dinate calibration events is selected by considering events located close to the tar-  
 372 get event and with magnitudes similar to the selected reference magnitude. We  
 373 assume that the reference magnitude is 2 and we select from the catalog events  
 374 having magnitude in the range 1.99-2.01 (we do not introduce errors in the nu-  
 375 merical analysis errors due to differences in location). The generated catalog con-  
 376 tains 202 candidates. Since we use  $Mw_{err}$  to perform the selection, the error on  
 377 magnitude contributes to the variance of the final residuals. Among the 202 candi-  
 378 dates, we randomly select 10 candidates and we average their spectra.
- 379 • ECS. For producing figures of intermediate results, we consider  $k = 0.025s$  (equa-  
 380 tion 8) and we assume  $\hat{f}_c=30$  Hz for the magnitude 2 gEGF. The 202 available  
 381 gEGFs, the 10 selected, and the average gEGF are shown in Figure S3 consider-  
 382 ing the scaling case A of Figure 7.

## 383 5.2 Results of the numerical test

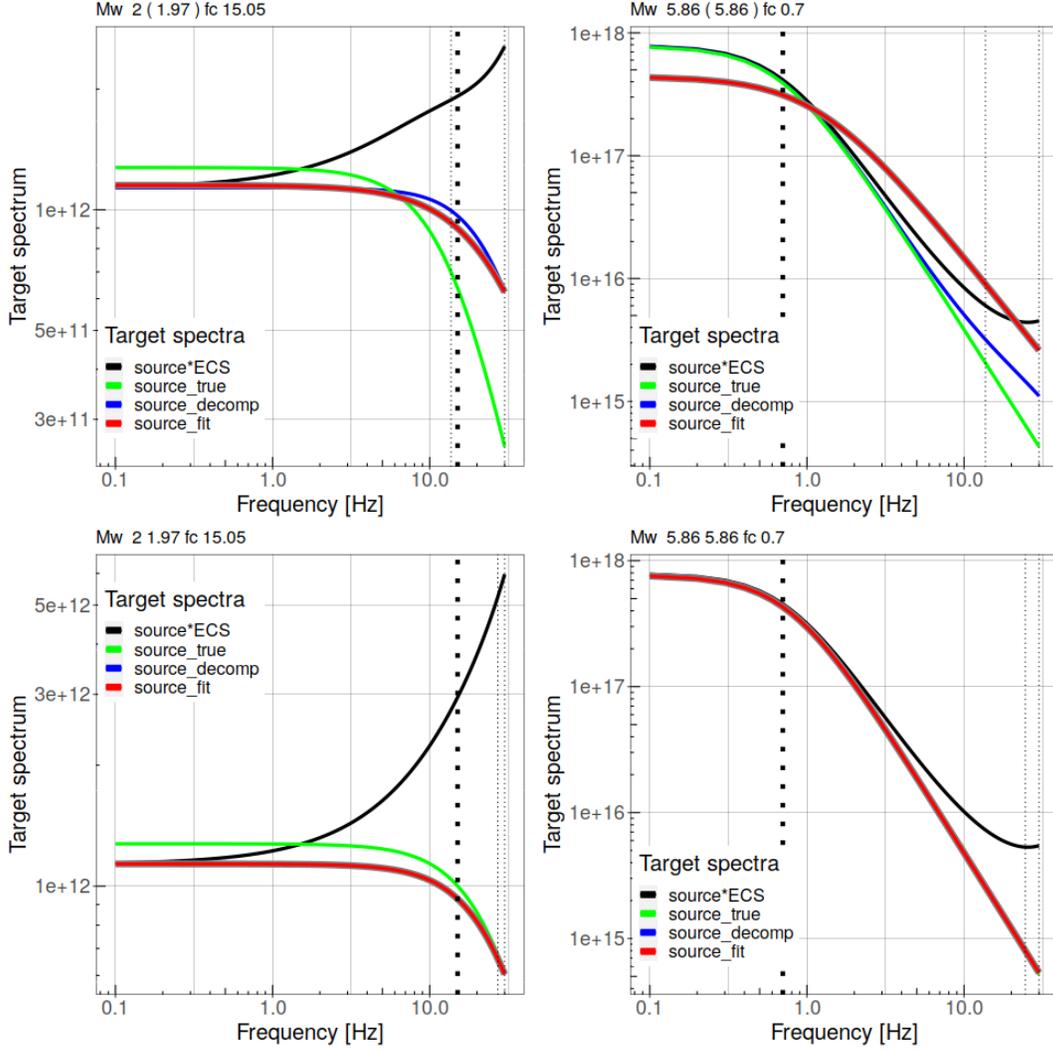
384 Figure 8 shows the results obtained considering the three different source scaling  
 385 of Figure 7. As expected, when the corner frequency assumed for the  $g\widehat{EGF}_{\hat{f}}$  deviates  
 386 from the true unknown value (as in the case of scaling A and B), the overall source scal-  
 387 ing is biased producing, in our case, corner frequencies higher than those used to gen-  
 388 erate the synthetic data. The overestimation is propagated also to magnitude larger than  
 389 the magnitude of the  $gEGF$ , as shown in Figure 8b: for magnitudes close to the  $gEGF$   
 390 magnitude, the overestimation is determined by the suitability of the  $g\widehat{EGF}_{\hat{f}}$  assump-  
 391 tion; for magnitudes larger than the  $gEGF$  magnitude, the overestimation decreases and  
 392 converge to an asymptotic average relative error of the order of 70 and 25 % for scaling  
 393 A and B, respectively. This apparent rotation of the overall source scaling caused by a  
 394 larger bias for small than for large events, impacts also the seismic moment. Whereas  
 395 the variability for a given magnitude is controlled by the error added to  $Mw$ , Figures  
 396 8d and 8f show that the impact of the wrong assumption on the  $g\widehat{EGF}_{\hat{f}}$  is propagated  
 397 to larger magnitudes, with scaling A and B generating an average  $Mo$  underestimation  
 398 above magnitude 5 of about 40 and 10 %, respectively. Figure 9 exemplifies the spec-  
 399 tral fitting for two synthetic events with magnitude 2 (left) and 5.9 (right). Results for  
 400 source scaling A are shown in the top panels, those for scaling C in the bottom ones. The  
 401 top left panel shows the worst case where the wrong assumption on  $f_c$  affects the  $g\widehat{EGF}_{\hat{f}}$   
 402 spectrum at frequencies close to the corner frequency of the source  $S$ . Moreover, the er-  
 403 rors added to  $Mw$  introduce an overall bias to the retrieved source spectrum  $S_{est}$ . In-  
 404 deed, the final fit is very good (small residuals) but with significant deviations on both  
 405  $f_c$  and  $Mo$  with respect to the true values. In other words, the precision on  $S_{est}$  is high  
 406 but its accuracy with respect to  $S$  is low. When scaling C is considered (bottom left panel),  
 407  $g\widehat{EGF}_{\hat{f}}$  is similar to the true  $gEGF$  and the high frequencies of the source spectra  $S$   
 408 are well retrieved. For magnitude 5.9 (right panels), even if the corner frequency of the source  
 409 is much smaller (0.7 Hz for the case shown in Figure 9) than  $f_c$  of  $g\widehat{EGF}_{\hat{f}}$ , Figure 8b  
 410 and 8d already showed that scaling A and B produce biased values for both  $Mo$  and  $f_c$ .  
 411 Whereas the added random error on  $Mw$  controls the variability of the  $Mo$  estimates,  
 412 the error on the  $g\widehat{EGF}_{\hat{f}}$  corner frequency has an impact on the overall shape of the source  
 413 spectrum, leading to biased estimates for both  $f_c$  and  $Mo$  (Figure 9, top right panel).  
 414 Figure 10 exemplifies the impact of the  $f_c$  assumption on the spectral shape consider-  
 415 ing two events with magnitude 3.14 and 4.08. The left panels show the results obtained  
 416 considering  $n = 2$  (omega square), i.e., the same value used for generating  $S$ ; in the right  
 417 panels, the fit performed by lowering  $n$  to 1.6 improves the quality of the fit being the  
 418 standard deviation of the residuals reduced by about a factor one-third (e.g., from 0.09147  
 419 to 0.03032 in the case of the magnitude 4.68 event). Therefore, the error on  $g\widehat{EGF}_{\hat{f}}$  led



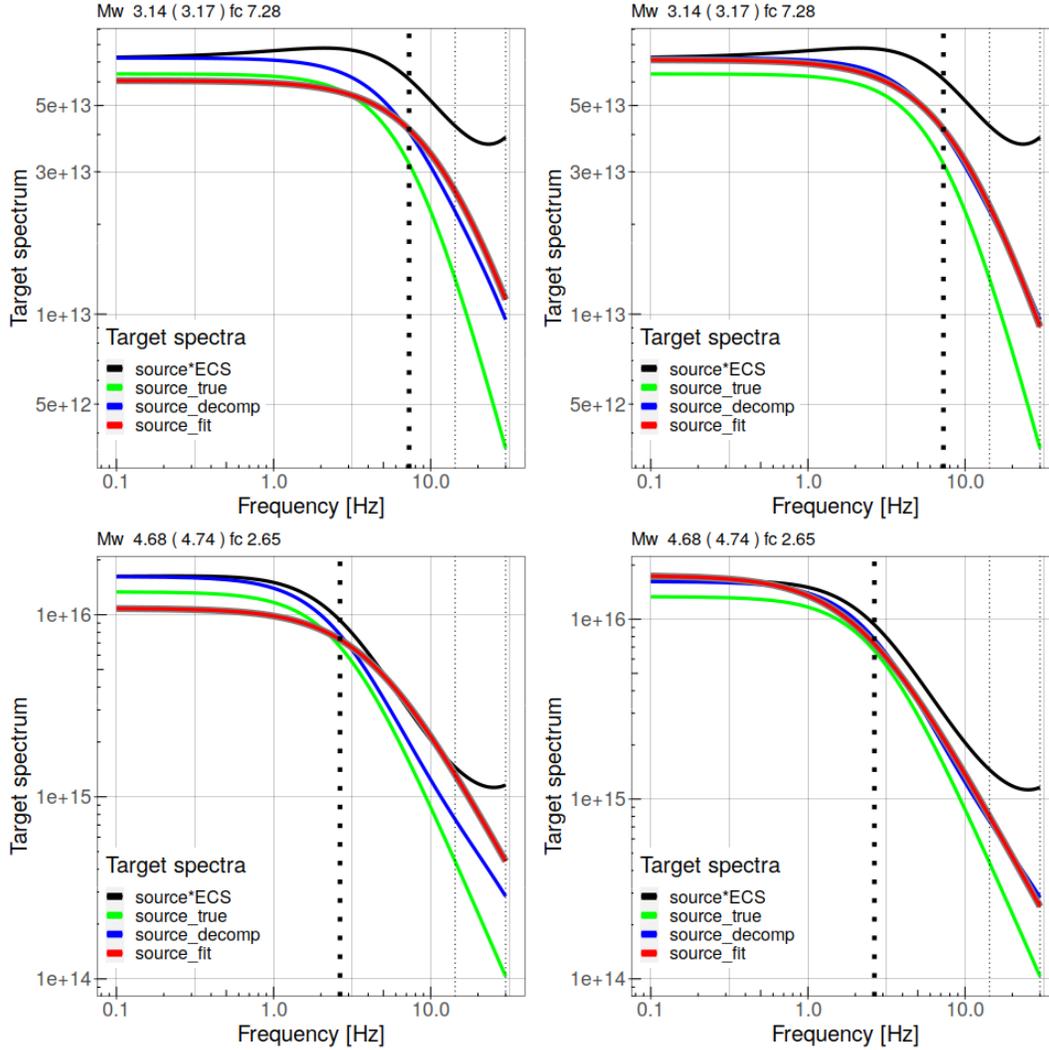
**Figure 7.** Source parameters used for the numerical test on the impact of constraining the corner frequency of small events. (a) uncertainty added to the magnitude values generated from the Gutenberg-Richter distribution in Figure ??; (b) assumed source parameters for the gEGF (cross) and three different stress drop scaling used for the numerical test (A is the most deviating scaling from the cross assumption; C is the most compatible scaling; B is the intermediate case); (c) variability added to the stress drop values for the same seismic moment; colors represent the three scaling in panel (b). (d) scaling accounting for the stress drop variability (panel c) and magnitude uncertainty (panel a).



**Figure 8.** Panels (a), (c) and (e) show the comparison between the retrieved scaling (white) and the simulated ones (colors represent scaling A, B and C as in Figure 7); results in terms of relative error for the corner frequency and seismic moment are shown in panels (b) and (d), respectively. Panel (f) shows the mean  $\pm 1$  standard deviation of the relative errors on seismic moment computed over different seismic moment bins.



**Figure 9.** Examples of the numerical test performed to evaluate the impact of constraining the corner frequency of small events. Results are exemplified in terms of spectral fitting considering a synthetic magnitude 2 (left) and a magnitude 5.9 (right) event. Top: results for the source scaling A of Figure 7 (worst case); bottom: results for the source scaling C of Figure 7 (best case). In the legend, source\*ECS indicates the numerator of equation 6, right hand side; source\_true is the true source spectrum  $S$ ; source\_decomposed is  $S_{est}$  (left hand side of equation 6); source\_fit is the final results, i.e., the best fit omega-square model of  $S_{est}$



**Figure 10.** Impact of considering  $n = 2$  (left) and  $n = 1.6$  (right). As in Figure 9, the true source spectrum  $S$  is indicated in legend as source.true, the estimated source spectrum  $S_{est}$  as source.decomp, the best fit source spectrum as source.fit, and the source\*ECS in black. Using  $n = 1.6$ , the quality of the fits improves although the correct source model used to generate the synthetics is  $n = 2$ .

to source spectra  $S_{est}$  with high-frequency slopes that deviate from the (true) omega-square model, allowing wrong models to fit the empirical source spectra better than models having the correct shape. This result confirms that we cannot discriminate among alternative source models by looking only at the quality of the fit.

## 6 Discussion

The main goal of this study was to recall some peculiarities of the non-parametric spectral decomposition approach and to discuss the consequences of strategies applied to select specific solutions. We developed our discussions by analyzing the characteristics of the design matrix compiled using a data set extracted from the Ridgecrest Stress Drop benchmark (Baltay et al., 2021). The most striking characteristic of the considered decomposition approach is the block-wise trade-offs among the source, propagation, and site terms, which results in a rank-deficient design-matrix with two null singular values. This characteristic is intrinsic to the decomposition approach, regardless of the number of available recordings or the model assumptions imposed on the unknown terms. By performing the singular value decomposition of the design matrix, we depicted the trade-off for the considered network geometry and earthquake catalog by looking at the singular vectors associated to null singular values. Since the rank deficiency implies the non-uniqueness of the solution (i.e., the matrix has a non-trivial kernel), solutions can only be selected relative to a-priori assumptions which generally vary from study to study. This is especially critical when comparing the results of different studies performed on the same data set, as in the case of the Ridgecrest stress drop benchmark. To avoid systematic differences due to the assumption of different reference conditions, it is necessary to consider procedures for comparing results in terms of relative quantities calculated within each study (Pennington et al., 2021), in addition to comparing absolute results between studies. Constraints to restore the uniqueness of the solution can be applied either to the source or to the site terms and, as a consequence of the block-wise trade-offs, they affect the unconstrained blocks as well. We have shown that the site constraint applied to the average amplification of selected reference stations corresponds to an average bias applied to the source spectra. Whereas information about site classification and empirical analysis can be used to design reliable site constraints, and to anchor the spectral decomposition models to site conditions of interest for engineering seismology applications, additional information on source characteristics are needed to assess, and possibly correct, biases transferred to the source. For example, seismic catalogs of moment magnitudes can be used as reference to anchor, on average, the seismic moments measured from the source spectra, by adding an overall vertical offset to the spectra (e.g., Bindi, Zaccarelli, & Kotha, 2020); the assumption of reference source spectral shapes (e.g., omega-square) can be used to evaluate the impact of the reference site condition at high frequencies by measuring the deviation from the expected acceleration spectral trend (e.g., Oth et al., 2011). Previous studies measured the deviation from the reference source model by coupling the omega-square model with a frequency dependent exponential term in which the exponent  $k_{source}$  parameter is controlling the high-frequency spectral decay. Investigations showed a good correlation between event-specific ground motion residuals and the  $k_{source}$  distribution (Bindi et al., 2019) suggesting that near source attenuation effects or differences in high-frequency radiation efficiencies contribute to the aleatory variability associated with ground motion prediction equations. Another common strategy for constraining the spectral decomposition solutions is to add a-priori information about the sources. To make solutions compatible with a-priori knowledge, source parameters of selected events (sometimes called reference events, ground truth events, or empirical Greens function) are imposed. Since parameters such as seismic moment, corner frequency, stress drop, and apparent stress are generally considered, spectral shapes and rupture models are also assumed, such as the omega-square and Brune models. We investigated the impact on the overall source scaling of constraining the corner frequency of small reference events by performing numerical simulations. Besides the

473 expected propagation of biases to corner frequencies and stress drops of small events when  
 474 the assumption deviates from the correct values, numerical tests showed that even the  
 475 source parameters of larger magnitudes are affected by the assumptions applied to small  
 476 events. Even when the corner frequencies of large events are small enough to fall far be-  
 477 low the spectral range affected by the approximations on the corner frequency of small  
 478 events, wrong assumptions for small events introduce a distortion of the source spectra  
 479 of large events. As a result, a fit performed using the correct spectral shape results in  
 480 source parameters different from those simulated. In addition, the quality of the fit from  
 481 the minimization of the residuals point of view can be improved by modifying the source  
 482 model. It follows that incorrect source models might be preferred to true models, sug-  
 483 gesting false deviations from the omega-square shape assumed for the simulations. Fi-  
 484 nally, it is worth mentioning the impact of model assumptions related to different terms  
 485 in the decomposition, particularly those concerning source and propagation. As shown  
 486 in equation 1, the spectral decomposition provides a single source spectrum correspond-  
 487 ing to the least-squares solution with respect to recordings available at several stations.  
 488 Since azimuthal dependencies are not taken into account, the retrieved spectra are as-  
 489 summed to represent spherically averaged quantities. Therefore, a good azimuthal and dis-  
 490 tance coverage is important to avoid systematic biases at low frequencies due to unbal-  
 491 anced sampling of radiation pattern effects and to avoid corner frequency biases due to  
 492 variability in source geometry, rupture directivity and rupture speeds (Kaneko & Shearer,  
 493 2015). Comparison of GIT source spectra with the distribution of the apparent source  
 494 spectra obtained by correcting the observations for GIT propagation and site effects can  
 495 provide useful indications for identifying possible azimuthal and distance-dependent ef-  
 496 fects (Bindi, Spallarossa, et al., 2020; Morasca et al., 2022). Because the back-scattering  
 497 effects that dominate the coda wavefield make coda analysis less sensitive to radiation  
 498 pattern effects (Aki, 1981; Mayeda & Malagnini, 2010), comparing GIT results with those  
 499 from coda is another effective strategy to quantify the impact of the spherical averag-  
 500 ing hypothesis. Regarding the propagation term, although few studies have implemented  
 501 more complex parametrizations for describing the attenuation term (e.g., Scherbaum,  
 502 1990; Edwards et al., 2008; Koulakov et al., 2010), it is generally assumed to be controlled  
 503 only by frequency and hypocentral distance. In the presence of significant lateral crustal  
 504 variability, unbalanced station distributions and clustered seismicity could favor the trans-  
 505 fer of attenuation effects to either the source or site terms. The analysis of the GIT resid-  
 506 uals can reveal the presence of spatial anomalies that could, for example, hint towards  
 507 the presence of areas with attenuation different from the average attenuation in the stud-  
 508 ied region. Considering the region in Figure 1, stronger than average attenuation can  
 509 be expected for the Coso volcanic area and weaker for the Mojave desert along the Gar-  
 510 lock fault (Hauksson & Shearer, 2006; Bindi et al., 2021, their Figure 10). Besides anisotropy  
 511 in the propagation medium characteristics, other phenomena not captured by the mod-  
 512 els used in equation 1 can generate coherent azimuthal patterns in the residuals, such  
 513 as directivity of the rupture. Recent investigations have also detected directivity effects  
 514 for small and moderate earthquakes (e.g., Gallovič, 2015; Pacor et al., 2016; Colavitti  
 515 et al., 2022) with significant modification of the source spectrum around the corner fre-  
 516 quency. Therefore, good station arrangement and high redundancy of the station and  
 517 event sampling are needed to mitigate the impact of the model assumption and misin-  
 518 terpretation. The characteristics of the data set contribute also to determine the prop-  
 519 erties of the covariance matrix. We have shown that the off-diagonal elements of the co-  
 520 variance matrix inform about trade-offs between existing between different couples of un-  
 521 knowns. For example, we showed that the trade-offs between attenuation and site am-  
 522 plification concentrate over the distance range sampled by the considered station. Again,  
 523 the quality of the data set in terms of recordings well distributed with distance and az-  
 524 imuth and high level of cross-information between stations and events are fundamental  
 525 to mitigate the impact of trade-offs generated by competing physical effects.

## 7 Conclusions

We investigated the non-uniqueness of the spectral decomposition approach applied to isolate source, propagation and site contribution to measured S-waves spectra. We related the non-uniqueness to the characteristics of the design matrix exemplifying the analysis for a data set extracted from the Community Stress Drop Validation Study (Baltay et al., 2021). We also discussed the impact on the final retrieved source scaling of constraints on site and the source terms applied to select specific solutions. We can summarize the main conclusions as follows:

- The problem of isolating source, propagation and site effects starting from a large compilation of S-wave amplitude Fourier spectra has a non-unique solution due to two unresolved degrees of freedom, regardless of the number of available recording, characteristics of the source catalog and network geometry; the non-uniqueness of the decomposition is intrinsic to any model framework where seismic measurements in either time or frequency domains are expressed as combination of source, propagation and site contributions;
- solutions can be only retrieved relative to a-priori constraints applied to break the numerical trade-offs affecting the three blocks (i.e., to fix the scale of the different blocks by assuming some reference anchoring points). Since the retrieved solutions are relative to these assumptions, comparison among results provided by different studies should be performed evaluating relative quantities (ratios or log differences) in order to remove the impact of the assumptions on the absolute values;
- Non-parametric spectral decompositions require data sets with high level of redundancy, i.e., each event recorded by several stations and each station recording several events; the covariance matrix allows to highlight the physical trade-offs between different unknowns, such as the site amplifications at any station and the attenuation coefficients for the distance range sampled by the station;
- Constraints applied to the site term have an impact on the source term; the selection of the reference stations and the imposed average amplification determines the average amplitude level of the source block: independent seismic moment values from a reference catalog should be used to correct for a possible bias. Moreover, the spectral shape of the reference amplification function has an impact on the average source spectral shape (e.g., on the average high frequency fall-off) that should be taken into account when the source spectra are fitted to a specific source spectral model.
- Numerical tests showed that assuming an a-priori value for the corner frequency (or stress drop) of small events (empirical Greens functions or reference events) can have a strong impact on the overall seismic moment versus corner frequency source scaling, introducing a bias for small events; the bias is also propagated to larger magnitudes but with reduced amplitude, generating an apparent rotation of the overall source scaling that could lead to wrong interpretations about the self-similarity of the rupture scaling. Moreover, the corner frequency assumption can modify the spectral shape of larger events, leading to best-fit source model shapes different from the correct one (e.g. with high frequency fall-off different from the values used for generating the synthetic data); consequently, alternative source models cannot be discriminated against solely on the basis of the quality of their fit.

In conclusion, we have shown that the well-known property of the decomposition approach of not having a unique solution can have a strong impact on the final interpretation of results, thus motivating our effort to raise awareness on this important topic for modelers and model users.

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## Open Research

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This work has been inspired by the Community Stress Drop Validation Study (Baltay et al., 2021) supported by the Southern California Earthquake Centre SCEC (<https://www.scec.org/research/stress-drop-validation>). The data set used in this study is disseminated through the benchmarks web-page. Routines for the simulations used for testing the source constraint have been prepared with R software (R Core Team, 2020) and are available on request.

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