

1 Reproductive dispersion and damping time scale 2 with life-history speed

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22 Abstract

23 Generation time has previously been the focus of comparative life history analyses. Here we
24 examine three metrics: generation time T_c , reproductive dispersion S (the distribution of
25 ages of reproduction), and damping time τ (time to converge to stable (st)age distribution).
26 We use data on 633 species of animals and plants, and perform phylogenetically corrected
27 analyses. First we find that S varies allometrically and isometrically with T_c . As a result, τ
28 varies allometrically with either T_c or S but not both. Second, we find a trade-off between τ
29 and S , so that τ does not vary isometrically with T_c . This trade-off is a novel demographic
30 component to the relationship between τ , T_c and S that is otherwise partly determined by
31 their similarity as biological times. Our results indicate that species at the slow end of the
32 slow-fast continuum take longer to converge to stable distribution than species with fast
33 life-histories.

34 Introduction

35 Life history traits describe processes such as survival, growth, and reproduction that deter-
36 mine an organism’s fitness components and are fundamental to ecological and evolutionary
37 processes, such as biological invasions (Sakai et al. 2001), local extinctions (Silvertown et al.
38 1996; Jongejans et al. 2008), and species diversification (Chesson 2000). Distinct combina-
39 tions of life history trait values result in distinct life history strategies (Stearns 1983), and
40 an important goal of life-history theory is to explain the range of variation in strategies
41 exhibited by species (Partridge and Harvey 1988; Hillesheim and Stearns 1992; Roff 1992,
42 2002; Salguero-Gómez et al. 2016b).

43 Given the diversity of life history patterns across species, it is notable that theory and
44 empirical evidence suggest that distinct life history strategies are placed along a slow-fast
45 continuum (Stearns 1983, 1992). Species at the slow end of the continuum are characterized
46 by late maturity, low fecundity and long lifespan, which lead to long generation time, and
47 those with the opposite suite of traits occupy the fast end and have short generation time
48 (Franco and Silvertown 1996; Oli 2004; Gaillard et al. 2016; Salguero-Gómez et al. 2016b;
49 Salguero-Gómez 2017). The “speed” of an age-structured life history can be measured by the
50 generation time T_c , the average age of net reproduction, where net reproduction is the prod-
51 uct of fertility $m(a)$ and the probability of survival $l(a)$ to age a . A similar generation time
52 can be constructed for stage-structured models (Cochran and Ellner 1992) and age+stage
53 structure (Steiner et al. 2014a). The generation time T_c correlates closely with the position of
54 a species along the slow-fast continuum (Gaillard et al. 2005) and is an important metric for
55 describing variation among life histories (Gamelon et al. 2014; Healy et al. 2019). Of course
56 the generation time does not reflect the age-spread of reproduction in iteroparous species,
57 which is measured by reproductive dispersion (also called demographic dispersion in Tul-
58 japurkar et al. (2009)). Certainly there is no connection in semelparous species that always
59 have zero reproductive dispersion regardless of generation time (Thomas 2013; Crespi and
60 Teo 2002) (note that semelparous individuals may be asynchronous in reproductive timing).
61 But in iteroparous species, the extent of this spread has long been of interest in life history
62 theory (Cole 1954; Trumbo 2013; Hughes 2017; Hautekèete et al. 2001; Varpe and Ejsmond
63 2018), and is our focus here. Let us quantify the reproductive dispersion by the standard
64 deviation S of the age of net reproduction around the average age T_c (other measures of re-
65 productive dispersion are considered later). The study of reproductive dispersion was made
66 famous by Cole (1954) who explored the difference between semelparity and iteroparity, and
67 there is ongoing interest in the extent of iteroparity (Hautekèete et al. 2001; Trumbo 2013;
68 Hughes 2017; Varpe and Ejsmond 2018). These studies suggest that the reproductive dis-
69 persion depends on many factors, including survival at juvenile and adult stages, trade-offs
70 between survival and reproduction, physiological development and regulation, environmental

71 variability, homeostatic ability, and behavior. In consequence we can argue that, in absence
72 of any structural constraints, there is no systematic covariation across a range of species
73 between the reproductive dispersion S and the average age of reproduction T_c .

74 However, structural constraints do matter (e.g. allometric constraints that cause all
75 life history traits expressed in mass, time, or length units to change with species-specific
76 body size)(McMahon and Bonner 1983), so we can make an alternative argument for such
77 covariation in iteroparous species. It is known that generation time scales with average
78 adult body mass (M) as $T_c \propto M^{0.25}$ (Millar and Zammuto 1983; Gaillard et al. 2005), like
79 all physiological times (Lindstedt and Calder 1981) or biological times (Lindstedt et al. 1986;
80 Gillooly et al. 2002; Brown et al. 2004; Hamilton et al. 2011). Given that both S and T_c are
81 “biological times”, i.e., internal, body-mass-dependent, time scales to which the durations
82 (or rates) of biological events are entrained (Lindstedt et al. 1986), we may expect that
83 reproductive dispersion scales in the same way with body size as generation time, and so
84 the two covary positively and isometrically sensu Huxley and Teissier (1936). But any such
85 relationship between S and T_c is likely to be noisy given the factors mentioned above.

86 A different perspective on reproductive dispersion comes from its effect on population
87 dynamics. To understand this, consider how a population structure returns to the stable
88 structure after a disturbance. After any such disturbance, reproduction has to fill in any gaps
89 in the population’s structure relative to the stable structure, and such gaps will be filled in
90 more rapidly in a species that has a high reproductive dispersion than in a species with low
91 reproductive dispersion. The time scale of population recovery after a disturbance is given
92 by the damping time. For age-structured populations (Keyfitz 1965; Coale 1972; Trussell
93 1977; Taylor 1979; Tuljapurkar 1982a,b, 1985; Wachter 1991; Caswell 2001), damping time
94 (τ) decreases with the reproductive dispersion (S) but increases with generation time (T_c) (a
95 similar relationship must hold for stage-structured populations). Previous analyses used this
96 result to explore the evolution of life histories (Orzack and Tuljapurkar 1989; Tuljapurkar
97 et al. 2009) but assumed that reproductive dispersion S and generation time T_c can vary
98 independently. While this might hold within species where evolutionary allometry is weak
99 at the best, large body size variation across species should lead to systematic covariation
100 between reproductive dispersion S and generation time T_c that will affect the pattern of
101 variation of damping time τ .

102 Hence, three hypotheses can be proposed about the association among generation time,
103 reproductive dispersion, and damping time across species that widely differ in size and
104 Baüplan. Under the biological time hypothesis (H1), all three metrics that are expressed
105 in time units should strongly covary positively and isometrically, leaving only a weak and
106 unstructured variation in each metric for a given life history speed. Under the demographic
107 hypothesis (H2), the damping time depends on both reproductive dispersion and generation

108 time. Lastly (hypothesis H3), both biological time and demographic hypotheses matter, so
109 that all three quantities should covary positively (in partial support of H1), but reproductive
110 dispersion and damping time should be partly decoupled from generation time. For a given
111 life history speed, a trade-off between reproduction dispersion and damping time should be
112 detected (in partial support of H2). Here we describe the results of a test of these hypotheses.

113 The next section defines reproductive dispersion, generation time, damping time, the
114 (known) analytical approximations, and the data we used. We then analyze the covariation
115 of reproductive dispersion and generation time, and of damping time. The results are based
116 on Phylogenetic Generalized Least Squares (PGLS, Freckleton et al. (2002)) regressions and
117 Phylogenetic Principal Component Analyses (PPCA, Revell (2010)) to account for the effect
118 of phylogeny. We also perform the analysis based on ordinary least squares regression (OLS)
119 and find similar results. The discussion considers the implications of our findings and the
120 possible reasons for them.

121 Definitions and Data

122 Reproduction, Dispersion and Damping

123 We use discrete times and ages. In an age-structured life history (Coale 1972; Caswell
124 2001; Keyfitz and Caswell 2005), with $m(a)$ and $l(a)$ as defined earlier, the expected lifetime
125 reproduction of a newborn is the net reproductive rate $R_0 = \sum_a l(a)m(a)$. For a cohort
126 (individuals born at the same time), the generation time (the measure we use for life history
127 “speed”) is the average age of net reproduction

$$T_c = \frac{\sum_a a l(a)m(a)}{R_0},$$

128 and the spread of reproduction around the mean age T_c is the reproductive dispersion S ,

$$S^2 = \frac{\sum_a (a - T_c)^2 l(a)m(a)}{R_0},$$

129 Similar definitions for stage-structure are given by Caswell (2001) and for age-and-stage
130 structure by Steiner et al. (2014b), and are used here. Note that there are alternate measures
131 for generation time such as T_b , calculated as weighted mean age of the mothers at childbirth
132 in a population (Gaillard et al. 2005). Our findings hold even if we use T_b in place of T_c
133 for generation time. Other measures of reproductive dispersion such as reproductive span
134 (the difference between age at last reproduction ω and age at first reproduction α) also yield
135 similar results as S . Additionally, when the population is stationary or nearly so, there is a

136 correlation between evolutionary entropy (Demetrius et al. 2009) and $\log S$.

137 The dynamics of a structured population are described by a population projection matrix
138 which has a dominant eigenvalue $\lambda_0 = \exp(r_0)$, where r_0 is the population intrinsic rate of
139 increase, and a leading subdominant eigenvalue $\lambda_1 = \exp(r_1 + is_1)$ (where r_1 , s_1 are the real
140 and imaginary parts and $i = \sqrt{-1}$). Here r_0 is always larger than r_1 (we assume that the
141 population matrix is irreducible and aperiodic). These eigenvalues define the damping time
142 τ as

$$\tau = \frac{1}{(r_0 - r_1)} > 0. \quad (1)$$

143 After a disturbance, the population structure approaches stability with time t as cycles
144 around the stable structure whose size (amplitude) decreases over time as $e^{-t/\tau}$. The am-
145 plitude of the cycles falls more slowly for a life history with high damping time than a life
146 history with small damping time. Thus the damping time in equation (1) is the time scale
147 of convergence of (st)age-structured populations to the stable (st)age distribution.

148 An approximate analysis by Wachter (1991) extended earlier work to show that

$$r_0 \approx \frac{\log R_0}{T_c} + \frac{S^2 (\log R_0)^2}{2 T_c^3}, \quad (2)$$

149 and

$$r_1 \approx r_0 - \frac{2\pi^2 S^2}{T_c^3}. \quad (3)$$

150 A similar approximation for r_0 holds in general for structured populations (Steiner et al.
151 2014b) and we conjecture that in such cases r_1 is similarly given by equation (3).

152 Using these approximations, the damping time in equation (1) is

$$\tau \simeq \frac{T_c^3}{2\pi^2 S^2}. \quad (4)$$

153 Hence, everything else being constant, damping time τ should increase with generation time
154 T_c , and decrease with increasing age dispersion S of reproduction. If the biological times T_c
155 and S scale perfectly with each other, then damping time τ will increase proportionally with
156 T_c , or alternatively with S .

157 We also compare the exact damping time τ , calculated from the population projection
158 matrix using equation (1) with the damping time given by the analytical approximation (4).
159 We find that the approximation does qualitatively predict the exact damping time on the
160 log-log scale (see Appendix B and Fig C.10).

161 Data

162 We calculate reproductive dispersion in a large number of species covering a wide range of
163 generation times. We use three databases which provide population projection matrices:
164 COMPADRE (v.6.20.5.0) for plants (Salguero-Gómez et al. 2015); COMADRE (v.4.20.5.0)
165 for animals (Salguero-Gómez et al. 2016a); and age-specific data for mammals compiled by
166 Jean-Michel Gaillard (Schindler et al. 2012) and Madan Oli (Oli 2004) (hereafter called GO).
167 These data have been uploaded as a supplement on *github*.

168 After data checking and cleaning (details in the Appendix A), we have a total of 3865
169 matrices (689 different species). To carry out phylogenetic generalized least squares (PGLS)
170 and phylogenetic principal component (PPCA) analyses, we begin with a master phylo-
171 genetic tree from Open Tree of Life version 12.3 ([https://tree.opentreeoflife.org/
172 about/synthesis-release/v12.3](https://tree.opentreeoflife.org/about/synthesis-release/v12.3)) and build a phylogenetic tree containing the species in
173 our analysis; we then use the *compute.brlen* function from the R package APE (Paradis and
174 Schliep 2019) to calculate the branch lengths. For species with multiple matrices, we use
175 the median value of reproductive dispersion S to select one matrix for each species as S
176 is more variable than T_c in the data (the mean of CV taken over all species is 0.43 for S
177 and 0.34 for T_c). The resulting dataset includes both phylogenies and life history traits for
178 633 out of 689 unique species (we could not find phylogenies from the master tree using
179 recorded scientific names for the other 56 species): stage-structured data for 319 species in
180 COMPADRE; stage-structured data for 215 species in COMADRE; age-structured data for
181 53 species in COMADRE; age-structured data for 75 species in GO. Thus we use these 633
182 species (633 matrices) for the PGLS analyses as well as PPCA. The data for plants are all
183 stage-structured, whereas animal data are structured by either stage or age (classification
184 discussed in the Appendix A).

185 In the graphs, results and discussions that follow, note that we standardize the time
186 unit as years. For each species, we compute directly (by standard numerical methods) the
187 life history speed (T_c), the reproductive dispersion (S), and the damping time (τ) from the
188 population projection matrix we retained.

189 We examine the data and our results in a variety of ways. In the main text, we present
190 PGLS and PPCA results separately for all age-specific data (from GO and COMADRE),
191 stage-specific data for animals (COMADRE), and stage-specific data on plants (COM-
192 PADRE). In the Appendix C, we present supplementary figures for PGLS regression for
193 each class of animals and plants.

194 Results

195 Generation time and reproductive dispersion

196 We first analyze the relationship between generation time (T_c) and reproductive dispersion
197 (S), computed as above from the datasets. As shown in Fig 1 for animals (top two panels),
198 and plants (bottom panel), $\log S$ is proportional to $\log T_c$. This conclusion holds within
199 animals and plants, and also within data organized by age-structure alone or stage-structure
200 alone. The relationship between $\log(S)$ and $\log(T_c)$ is strong ($R^2 > 0.8$) and statistically
201 significant ($P \leq 0.001$). The regression of $\log(S)$ versus $\log(T_c)$ in Fig 1 yields slopes between
202 0.96 and 1.12, all close to the value of 1 that corresponds to isometry. Pagel’s λ for animals
203 and plants structured by stage is 0.29 and 0.00, respectively, which indicates that phylogeny
204 has only a weak influence. For animals structured by age, Pagel’s λ is 0.63, indicating a
205 moderate influence of phylogeny.

206 Does a similar relationship between reproductive dispersion and generation time hold for
207 the species within taxonomic classes? We group species by classes, and keep those classes
208 that contain 20 or more species, including *Actinopterygii*, *Aves*, *Mammalia*, and *Reptilia* in
209 animals and *Magnoliopsida* and *Liliopsida* in plants. A strong scaling relationship between
210 reproductive dispersion S and generation time T_c is found in each class. The regression slope
211 between $\log T_c$ and $\log S$ for each class varies but is approximately close to 1 (isometry).
212 Summary statistics for each class such as Pagel’s λ and R^2 are reported in the Appendix C
213 (Figs C.1, C.2).

214 For just age-structured populations we can argue that an increase in reproductive dis-
215 persion likely implies a large reproductive span as measured by the difference between age
216 at last reproduction ω and age at first reproduction α . In such cases, the biological time
217 hypothesis implies an allometric and isometric increase in reproductive span ($\omega - \alpha$) with
218 generation time, and indeed we find such a relationship, although noisy (see Appendix Fig
219 C.3).

Figure 1 here

220 Three time metrics: damping, generation, reproductive dispersion

221 We next examine how the damping time τ (calculated from the population projection matrix)
222 changes with generation time T_c and reproductive dispersion S , for age or stage structure.

223 Fig 2 for animals and plants shows that $\log \tau$ is proportional to $\log T_c$ even though the
224 relationship is not isometric. Across species we find that

$$\tau \propto T_c^b, \text{ with } b \text{ between } 0.65 \text{ and } 0.82. \quad (5)$$

225 The relationship between $\log(\tau)$ and $\log(T_c)$ is statistically significant but noisy, more so for
226 stage-based population models than age-based ones. In addition, the positive relationship
227 holds within classes (see Appendix Fig C.4 and C.5), but is still noisy.

228 The variability around the main correlation in Fig 2 could be due to the effect of repro-
229 ductive dispersion, in addition to the inevitable variation due to the effect of sample size
230 or the rarity of some events (e.g., such as deaths of large trees). However, we find that the
231 variability in Fig 2 is largely independent of reproductive dispersion (S) (see Appendix Fig
232 C.6).

233 Given our previous finding that S scales allometrically with T_c (see comments following
234 equation (4)), we expect that $\log \tau$ is also proportional to $\log S$. This is indeed what we find
235 (see Appendix Fig C.7). These results clearly show that the demographic hypothesis (H2)
236 is false, but only partly support the biological time hypothesis (H1). Although damping
237 time and generation time are positively associated, they are linked with a hypoallometric
238 relationship, meaning that the time to converge to stable (st)age structure increases with
239 generation time but more slowly than if there was the simple proportionality expected under
240 H1.

Figure 2 here

241 The strong covariation among generation time, reproductive dispersion, and damping
242 time means that we must contend with multicollinearity (Dormann et al. 2013). We therefore
243 turn to a PPCA on the three datasets (i.e. age-structured animals, stage-structured animals,
244 and stage-structured plants). Fig 3 shows the first two principal components (PPC1 and
245 PPC2) for age-structured animals. In the Appendix C we show the corresponding results in
246 Fig C.8 and C.9 for stage-structured animals and plants. Details of the PPCA are presented
247 in Table 1. The results were remarkably consistent across these datasets indicating that the
248 nature of population structure (i.e. age vs. stage) or the type of organism (animal vs. plant)
249 does not influence life history patterns.

250 As expected under H1, PPC1 captures the positive covariation among the three metrics
251 and accounts for most across-species life history variation. However, this outcome does not
252 fully support H1. First, PPC1 only explains around three-quarter of the total life history
253 variation included in the three metrics (74% for age- and stage-structured animals, 77% for
254 plants). Second, generation time by itself almost entirely determines the time scale corre-
255 sponding to PPC1 (86% for age-structured animals, 87% for stage-structured animals, and
256 98% for plants), so we see that PPC1 is aligned with the slow-fast continuum. Both repro-
257 ductive dispersion and damping time show a substantial variation that is decoupled from the
258 slow-fast continuum, so PPC2 emerges as a second structuring axis of life history variation.
259 This second axis accounts for substantial life history variation (23% for age-structured ani-
260 mals, 26% for stage-structured animals, and 22% for plants) and leaves the third potential

261 axis to be restricted to noise by accounting only for 3% or less of life history variation (3% for
262 age-structured animals, 0.6% for stage-structured animals, 0.7% for plants). PPC2 describes
263 the trade-off expected under H3 between reproductive dispersion and damping time. Plants
264 and animals differ in the relative contribution to this trade-off between reproductive disper-
265 sion and damping time. In animals, reproductive dispersion contributes only slightly more
266 than damping time to PPC2 (loadings for age-structure -0.61 vs. 0.49, for stage-structure
267 -0.57 vs. 0.54), consistently lower than the contribution of these traits to PPC1. In plants,
268 damping time contributes much more than reproductive dispersion to PPC2 (-0.24 vs. 0.85),
269 with a relative contribution much larger than to PPC1.

Figure 3 here

Table 1 here

270 Discussion

271 Our first finding here is a robust scaling relationship between reproductive dispersion S
272 and generation time T_c . Across a wide range of species and taxa, S varies allometrically
273 and approximately isometrically with T_c , in partial support of hypothesis H1. This finding
274 implies that species with slow life-histories (characterized by late maturity, low fecundity,
275 and long generation time) typically spread reproduction over a wider age range than those
276 with fast life-histories. A population that has both relatively high T_c and dispersion S
277 is likely better able to time-average risks (e.g., during development, during reproduction).
278 Such bet-hedging strategies may help ride out fluctuations by dispersing the effect over large
279 (st)age classes and have been discussed by Tuljapurkar et al. (2009) and Sæther et al. (2013).

280 Another consequence of the biological time hypothesis (H1) is that damping time τ
281 should vary allometrically and isometrically with generation time T_c . But we only find that
282 τ is positively correlated with T_c on the log-log scale, with modest effects of reproductive
283 dispersion S . The slope of $\log \tau$ on $\log T_c$ is consistently less than 1, and so the time to
284 converge to stable (st)age structure has a hypoallometric scaling with generation time, which
285 contrasts with the isometric relationship expected under H1.

286 Nonetheless, this finding implies that species with slow life-histories take a long time
287 to converge to stable structure after disturbances than species with fast life-histories. Our
288 finding greatly extends Capdevila et al. (2020) who argue that the long convergence time
289 (slow recovery rate) of the Asian elephants (*Elephas maximus*) makes them more vulnerable
290 to continuous habitat loss than red squirrels (*Tamiasciurus hudsonicus*). We propose that
291 generation time, a major axis of variation in mammalian life-history tactics (Gaillard et al.
292 2005), also sets the time scale for response and recovery of species following perturbations.

293 Alternatively, given that we have found that T_c and S are strongly correlated, we could

294 state that τ covaries with S on the log-log scale, with small contribution from T_c . Indeed,
295 we find such a relationship, and so hypothesis H2 is rejected because of the covariation of T_c
296 and S . To deal with the collinearity of T_c and S , we use a PPCA.

297 Our PPCA results provide clear evidence of selective pressure to reduce damping time.
298 The first principal component (PPC1) aligns with the slow-fast continuum. However, we
299 find that PPC2 accounts for about one quarter of overall variation among the three traits
300 analyzed, independent from the slow-fast continuum (PPC1). PPC2 is largely shaped by a
301 trade-off between reproductive dispersion and damping time. This trade-off is pervasive and
302 occurs independently of the population structure or of organisms considered. This trade-off
303 is also consistent with our last hypothesis (H3) which corresponds to a general demographic
304 process that has remained undetected up to now. However, selective pressure against long
305 damping time seems to differ between plants and animals. In animals, reproductive disper-
306 sion plays a key role by being weakly coupled to the slow-fast continuum. On the other
307 hand, in plants, damping time is much more strongly decoupled from the slow-fast contin-
308 uum. We thus propose that in animals an increase in reproductive dispersion limits the long
309 damping time that might accompany a long generation time, whereas in plants an increase
310 of reproductive dispersion has a smaller effect on damping time.

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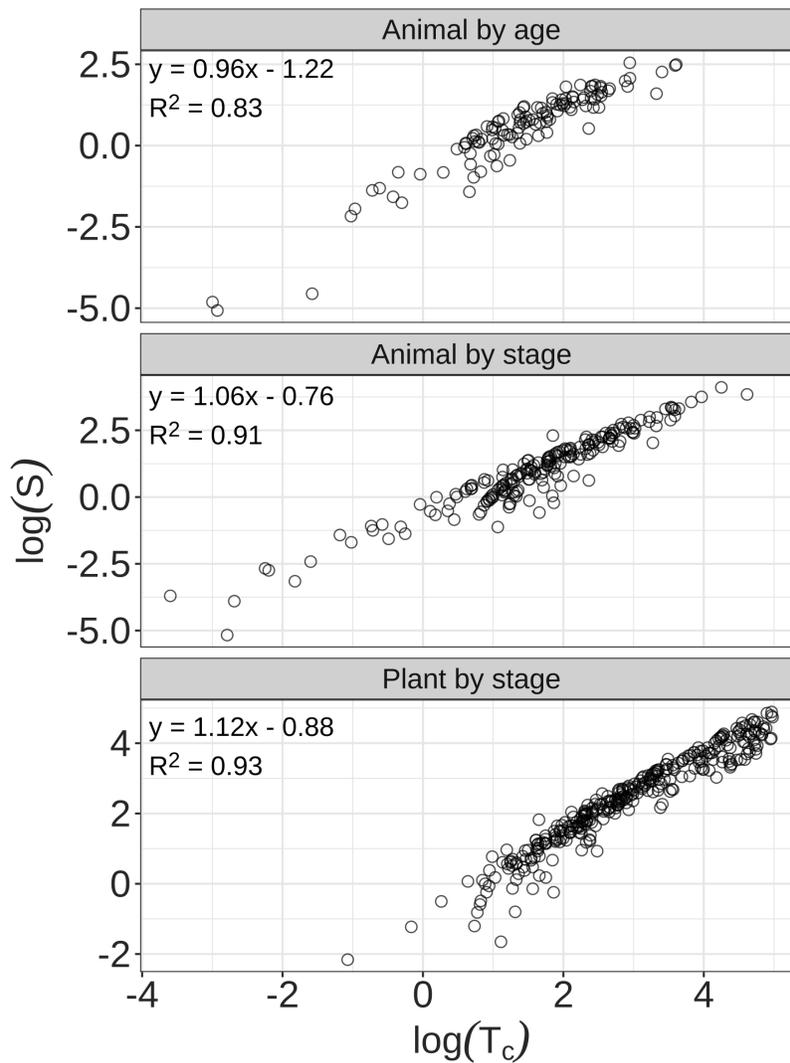


Figure 1: Reproductive dispersion (S) versus generation time (T_c) on a log-log scale. Time unit is years. Upper panel, age-structured animal data from COMADRE and GO; middle panel, stage-structured animal data from COMADRE; bottom panel, stage-structured plant data from COMPADRE. Each panel displays the fitted model and its coefficient of determination (R^2) based on PGLS regression. P-value in each panel is less than 0.001. The 95% confidence interval for the regression slope is [0.92, 1.00] for the upper panel, [1.04, 1.09] for the middle panel and [1.11, 1.14] for the bottom panel. Pagel's λ is 0.63, 0.29 and 0.00 for the upper, middle and bottom panels respectively.

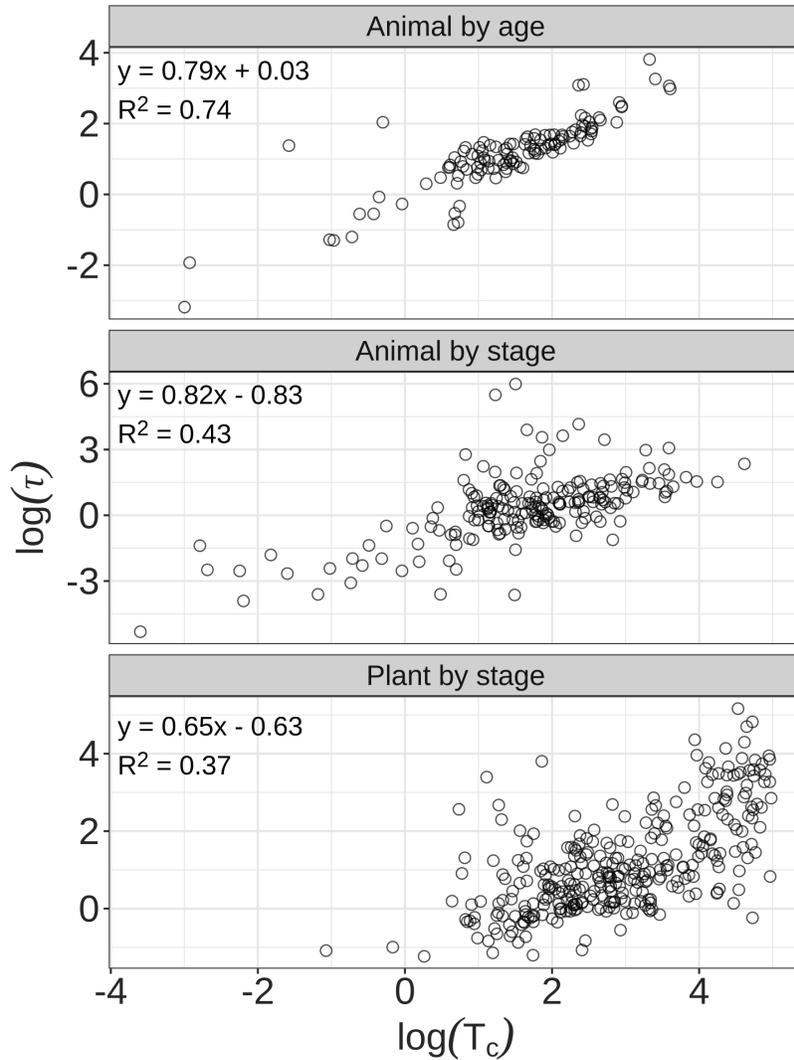


Figure 2: Damping time (τ) versus generation time (T_c) on a log-log scale. Damping time (τ) is calculated directly from each population projection matrix. Time unit is years. Upper panel, age-structured animal data from COMADRE (circles) and GO (crosses); middle panel, stage-structured animal data from COMADRE; bottom panel, stage-structured plant data from COMPADRE. Each panel displays the fitted model and its coefficient of determination (R^2) based on PGLS regression. P-value in each panel is less than 0.001. The 95% confidence interval for the regression slope of each panel is [0.74, 0.83] for the upper panel, [0.75, 0.89] for the middle panel and [0.61, 0.70] for the bottom panel. Pagel's λ is 0.00, 0.13 and 0.27 for the upper, middle and bottom panels respectively.

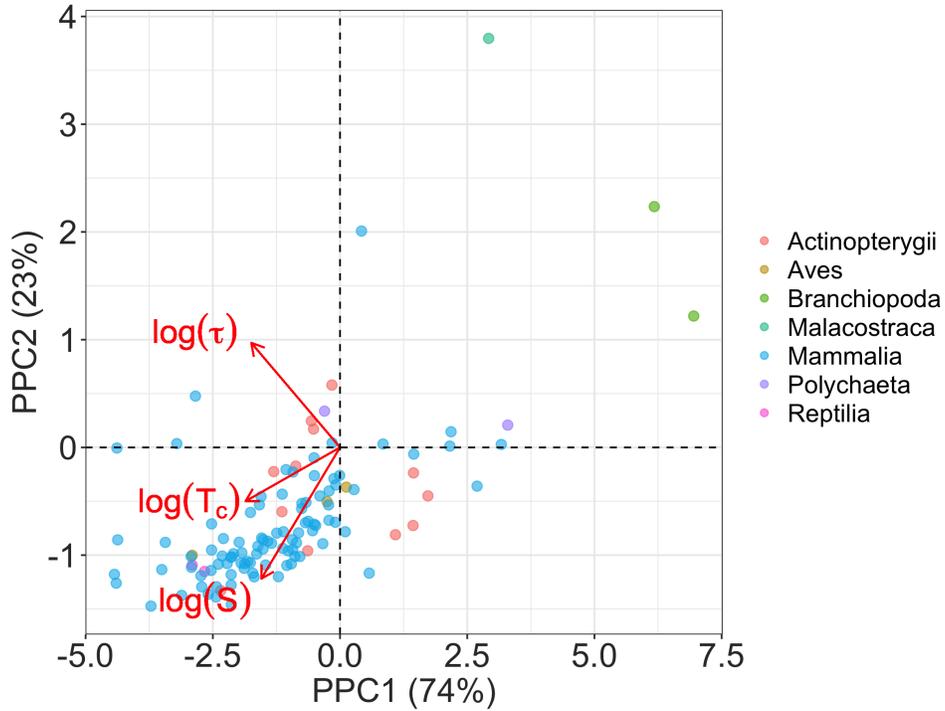


Figure 3: Phylogenetic principal component analysis (PPCA) of reproductive dispersion (S), generation time (T_c) and damping time (τ) on a log scale for age-structured animals. Arrow length indicates the loading of each life-history trait onto PCA axes. Points represent the position of species along the phylogenetically corrected principal component (PPC)1 and 2, and are colored by Class. Numbers in parentheses on both axes represent the proportion of variance explained by the corresponding PPC. Results for stage-structured animals and plants can be found in the Appendix C.

Table 1: Loadings from a phylogenetic principal component analysis (PPCA)

Dataset		Principal component	
		PPC1	PPC2
Animal by age	Metric		
	$\log(T_c)$	-0.93	-0.25
	$\log(S)$	-0.77	-0.61
	$\log(\tau)$	-0.87	0.49
	Proportion of variance	74%	23%
	Cumulative proportion	74%	97%
Animal by stage	Metric		
	$\log(T_c)$	0.93	-0.34
	$\log(S)$	0.82	-0.57
	$\log(\tau)$	0.84	0.54
	Proportion of variance	74%	26%
	Cumulative proportion	74%	99%
Plant by stage	Metric		
	$\log(T_c)$	-0.99	-0.07
	$\log(S)$	-0.97	-0.24
	$\log(\tau)$	-0.53	0.85
	Proportion of variance	77%	22%
	Cumulative proportion	77%	99%

447 **Appendix**448 **A Supplementary Information on Data**

449 For COMPADRE (v.6.20.5.0) and COMADRE (v.4.20.5.0) database, there are initially 8925
450 matrices (759 species) and 2275 matrices (415 species), respectively. Then we conduct a series
451 of data cleaning to prepare the dataset for the analysis.

452 **A.1 Data classification for age- and stage-structured matrices**

453 The COMADRE and COMPADRE database consists of three criteria to indicate whether
454 the population projection matrix contains (st)ages based on size (MatrixCriteriaSize), devel-

455 opment (MatrixCriteriaOntogeny), age (MatrixCriteriaAge). We get a rough classification
456 of age and stage-structured data after removing 'NA' in the three criteria. If MatrixCrite-
457 riaSize == "No" & MatrixCriteriaAge == "Yes" & MatrixCriteriaOntogeny == "No", it's
458 considered as age-structured data, otherwise it's stage-structured data.

459 For COMPADRE, we only consider stage-structured data for the analysis. For age-
460 structured data in COMADRE, we further check the intersection of last row and last column
461 in the survival matrix. If the value is zero, then we classify it as a age-structured data; if it
462 is non-zero, then classify it as stage-structured data considering that some individuals will
463 survive beyond the last age observed.

464 A.2 Filters used before calculation

465 Using the flags in the dataset, we exclude data with missing values in vital rates (i.e, no
466 NA's in the matrices); ensure the survival probability is always less than 1 for the last
467 (st)age and less than or equal to 1 for other (st)ages; ensure the survival probability for
468 the last age/stage is always less than 1; ensure the fecundity was measured in the study;
469 remove data from one unclear source (Master thesis with no title and author name); remove
470 semelparous species *Oncorhynchus tshawytscha* (Chinook salmon), Bacteria (*Spirochaetes*)
471 and Virus (*lentivirus*).

472 We keep those fertility matrices that have non-zero elements only on the first row to ensure
473 offspring are born into the first (st)age. We eliminate those matrices with non-zero cloning
474 data since we do not analyse clonal mode of reproduction. To standardize the life-history
475 traits to units of year based on the values of projection intervals, we only keep matrices where
476 the projection interval is non-zero. For data that have mixed male and female population
477 projection matrices in one matrix, marked as StudiedSex == "M/F - Males and females
478 separately in the same population matrix model" in the database, we separate and check
479 them individually according to their source papers.

480 For age-structured data in COMADRE, we further ensure survival matrix should have
481 non-zero value only in the sub-diagonal; ensure that the fertility matrix has more than 1
482 non-zero value in the first row to allow for dispersion of reproductive events.

483 A.3 Filters used during and after calculation

484 For stage-structured data in both COMADRE and COMPADRE, we remove matrices where
485 (I-U) inverse does not exist (where I is the Identity matrix and U is the survival matrix)
486 to enable the calculation. Besides, considering the biological realisticity, we remove unlikely
487 values by ensuring $\log(T_c) < 5$, $\log(S) > -15$, and $\log(\tau) < 15$.

488 A.4 Matrices used for PGLS analysis

489 We then combine our computed life-history traits (generation time, reproductive dispersion
490 and damping time) with the phylogenetic tree we build based on the master phylogenetic
491 tree from Open Tree of Life version 12.3 ([https://tree.opentreeoflife.org/about/synthesis-](https://tree.opentreeoflife.org/about/synthesis-release/v12.3)
492 [release/v12.3](https://tree.opentreeoflife.org/about/synthesis-release/v12.3)). For several species we had multiple matrices, we use the median value of
493 age-dispersion S to select one matrix for each species as S was much more variable than T_c
494 in the data (the mean of CV within each species is 0.43 for S and 0.34 for T_c). The resulting
495 dataset includes both phylogenies and life history traits for 633 out of 689 unique species
496 (we could not find phylogenies from the master tree using recorded scientific names for the
497 other 56 species). Thus we used these 633 species (633 matrices) for the PGLS analyses:
498 stage-structured data for 319 species in COMPADRE; stage-structured data for 215 species
499 in COMADRE; age-structured data for 53 species in COMADRE; age-structured data for
500 75 species in GO.

501 A.5 Limitations

502 Small population sizes may lead to biased estimates of vital rates. In some populations that
503 have long-lived stage(s), such as trees, the numbers of deaths to large individuals observed
504 during the study period may be small so the corresponding estimated survival rates may be
505 artificially high. Consequently, studies that incorporate both age and stage structure will
506 be important (de Valpine et al. 2014). Matrix dimensionality is taken from the data, but is
507 known to influence the life-history traits calculated (Salguero-Gomez and Plotkin 2010).

508 B How Good is the Analytical Approximation?

509 Recall that equation (4) approximates damping time τ in terms of the ratio T_c^3/S^2 . To eval-
510 uate this approximation, we compare the exact damping time τ (calculated directly from the
511 population projection matrix) and the damping time τ given by the analytical approxima-
512 tion. On a log-log scale, we find that the analytical approximation does qualitatively predict
513 the exact damping time, indicating a statistically significant correlation between them (see
514 Appendix Fig C.10). In general, the damping time from the approximation is smaller than
515 the damping time calculated from data. Compared to plants, the analytical approximation
516 for animals is better because the OLS regression for animals has a slope close to 1 (0.94 for
517 age-structured animals, 0.84 for stage-structured animals, and 0.63 for plants) and a large
518 R^2 (0.71 for age-structured animals, 0.79 for stage-structured animals, and 0.54 for plants).
519 This could be because the approximation ignores higher moments of the distribution of
520 reproduction, which may be significant for plants and animals with stage-based dynamics.

C Figure

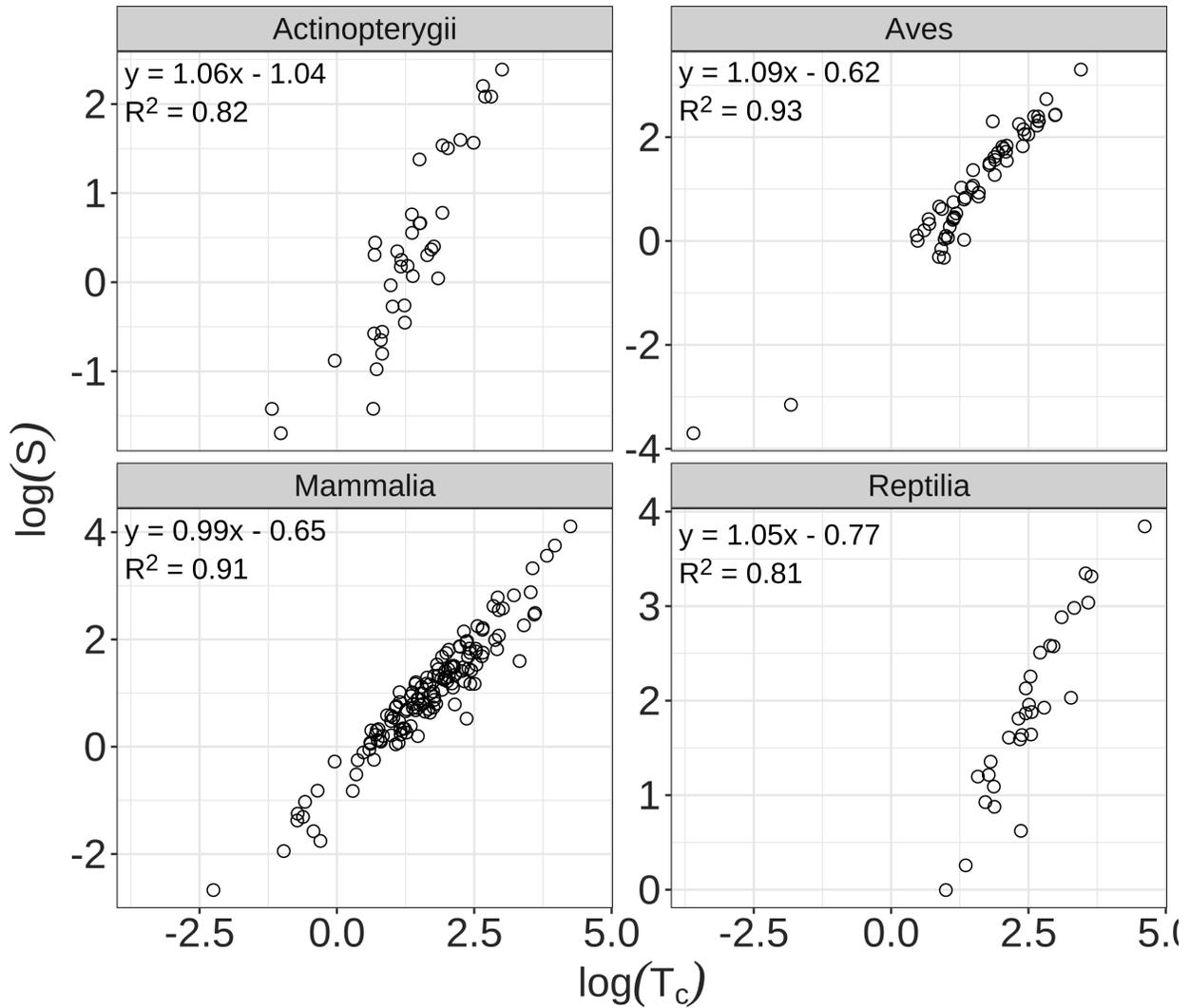


Figure C.1: Class-wise plots for reproductive dispersion (S) versus generation time (T_c) on a log scale for animals. The unit of time is years. Each panel corresponds to a Class. On the top left of each panel, we also present the fitted model and its coefficient of determination (R^2) based on PGLS regression. P-value in each panel is less than 0.001. The 95% confidence interval for the regression slope and the value of Pagel's λ are [0.98, 1.14] and 0.50 for *Actinopterygii*, [1.05, 1.13] and 0.00 for *Aves*, [0.96, 1.01] and 0.01 for *Mammalia*, [0.95, 1.14] and 0.01 for *Reptilia*.

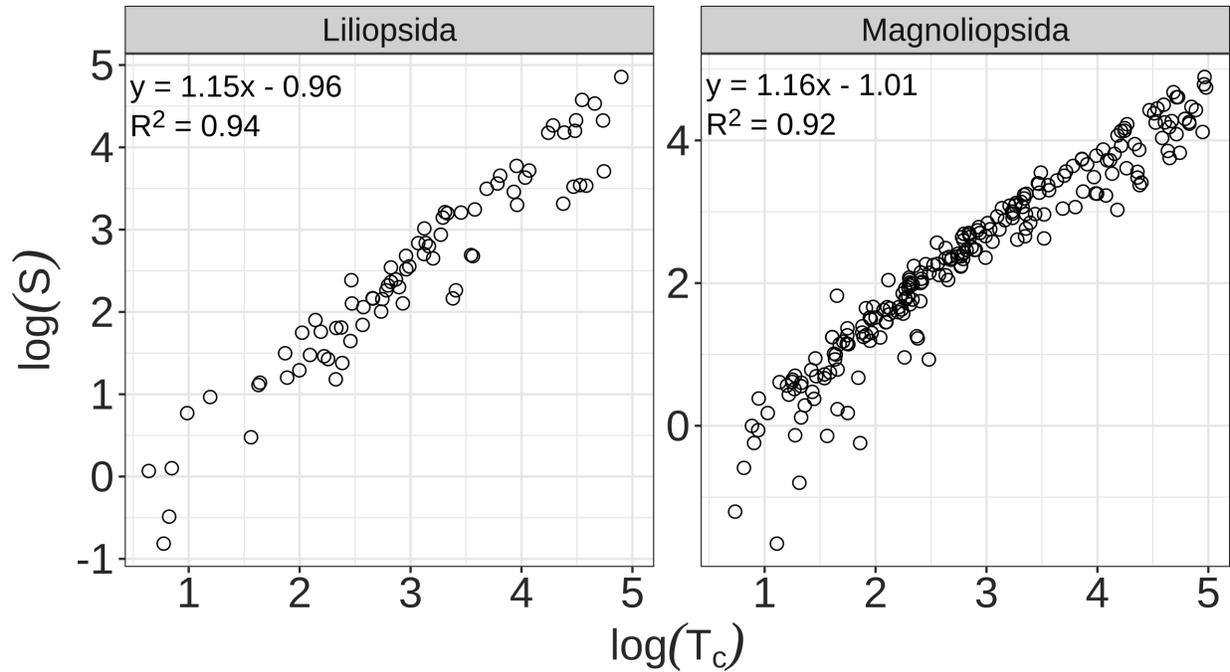


Figure C.2: Class-wise plots for reproductive dispersion S versus generation time T_c on a log scale for plants. The unit of time is years. Each panel corresponds to a Class. On the top left of each panel, we also present the fitted model and its coefficient of determination (R^2) based on PGLS regression. P-value in each panel is less than 0.001. The 95% confidence interval for the regression slope and the value of Pagel's λ are [1.11, 1.18] and 0.83 for *Liliopsida*, [1.13, 1.18] and 0.05 for *Magnoliopsida*.

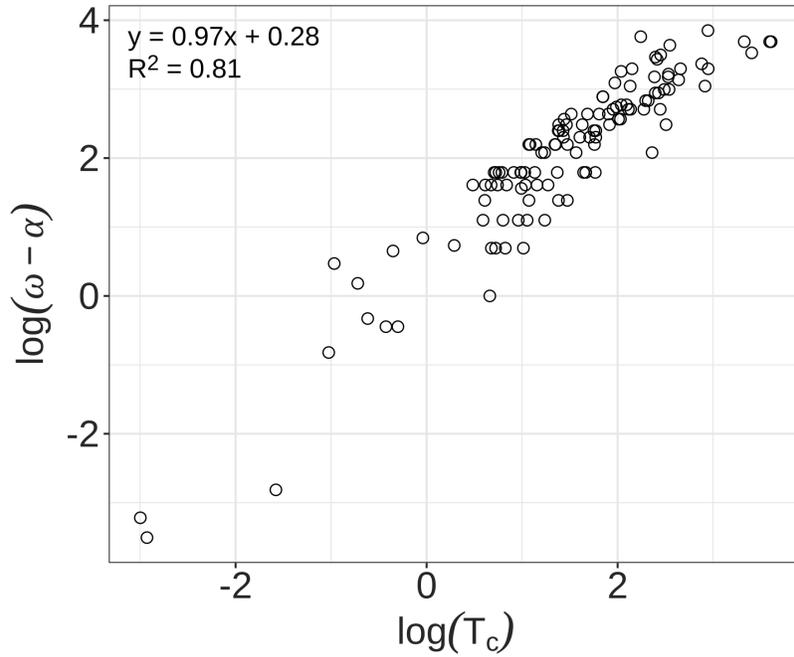


Figure C.3: Reproductive span ($\omega - \alpha$) versus generation time (T_c) on a log scale for age-structured animals. The unit of time is years. On the top left, we also present the fitted model and its coefficient of determination (R^2) based PGLS regression. P-value is less than 0.01. The 95% confidence interval for the regression slope and the value of Pagel's λ are [0.92, 1.01] and 0.60.

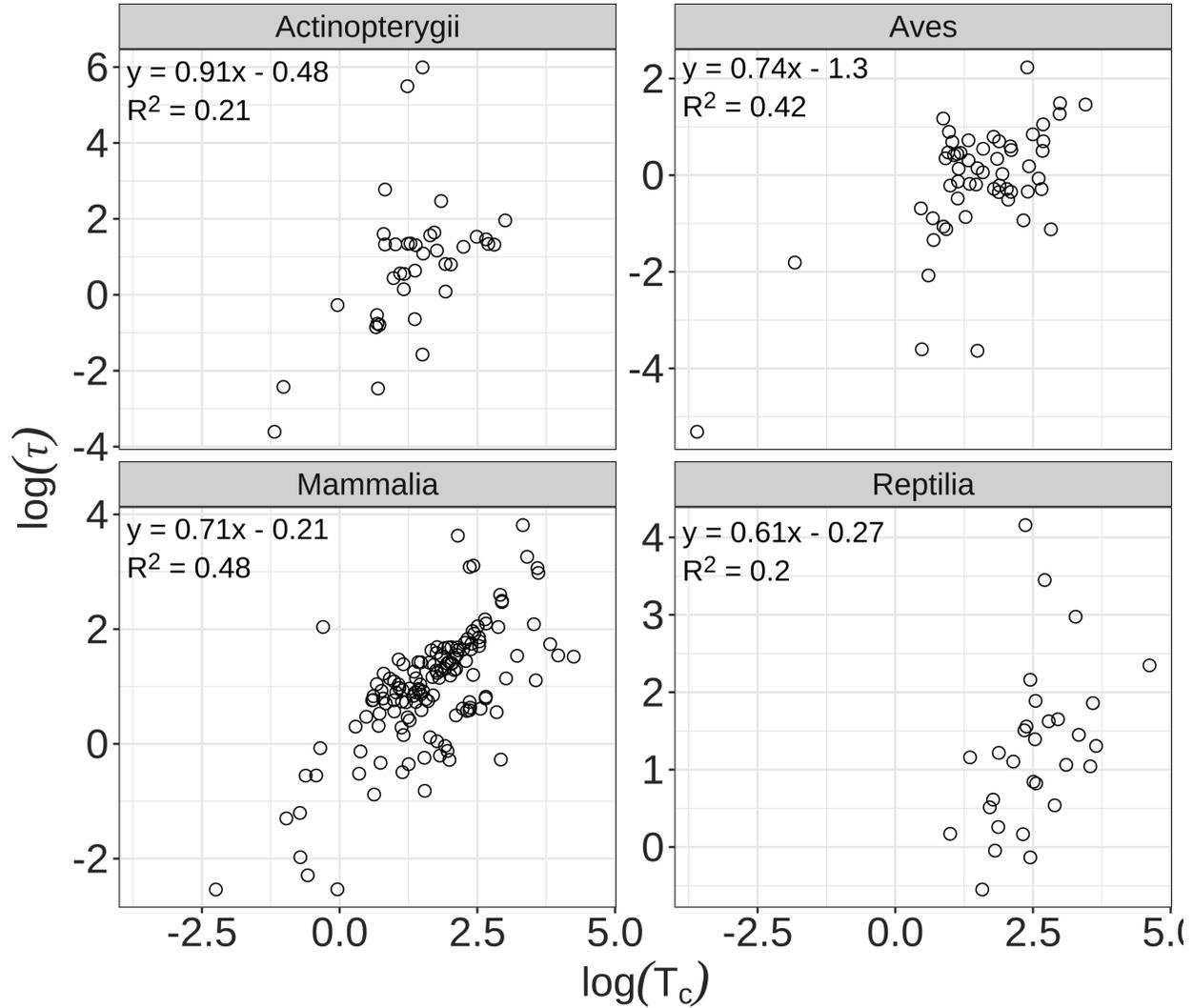


Figure C.4: Class-wise plots for damping time (τ) versus generation time (T_c) on a log scale for animals. It should be noted that the damping time (τ) presented here is the exact value calculated from population projection matrix instead of the approximation in equation (4). The unit of time is years. Each panel corresponds to a Class. On the top left of each panel, we also present the fitted model and its coefficient of determination (R^2) based on PGLS regression. P-value in each panel is less than 0.01, except for *Reptilia* ($P = 0.01$). The 95% confidence interval for the regression slope and the value of Pagel's λ are [0.62, 1.20] and 0.52 for *Actinopterygii*, [0.62, 0.86] and 0.00 for *Aves*, [0.65, 0.77] and 0.00 for *Mammalia*, [0.38, 0.84] and 0.00 for *Reptilia*.

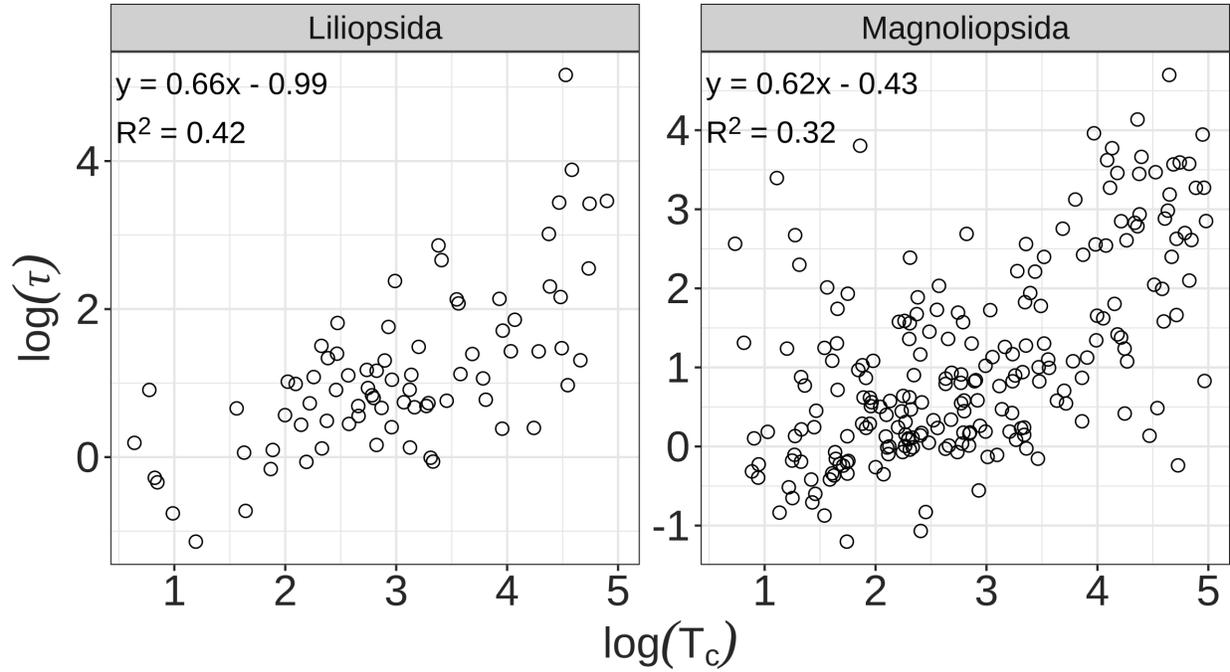


Figure C.5: Class-wise plots for damping time (τ) versus generation time (T_c) on a log scale for plants. The unit of time is years. Each panel corresponds to a Class. It should be noted that the damping time (τ) presented here is the exact value calculated from population projection matrix instead of the approximation in equation (4). On the top left of each panel, we also present the fitted model and its coefficient of determination (R^2) based on PGLS regression. P-value in each panel is less than 0.01. The 95% confidence interval for the regression slope and the value of Pagel's λ are [0.57, 0.74] and 0.45 for *Liliopsida*, [0.56, 0.68] and 0.12 for *Magnoliopsida*.

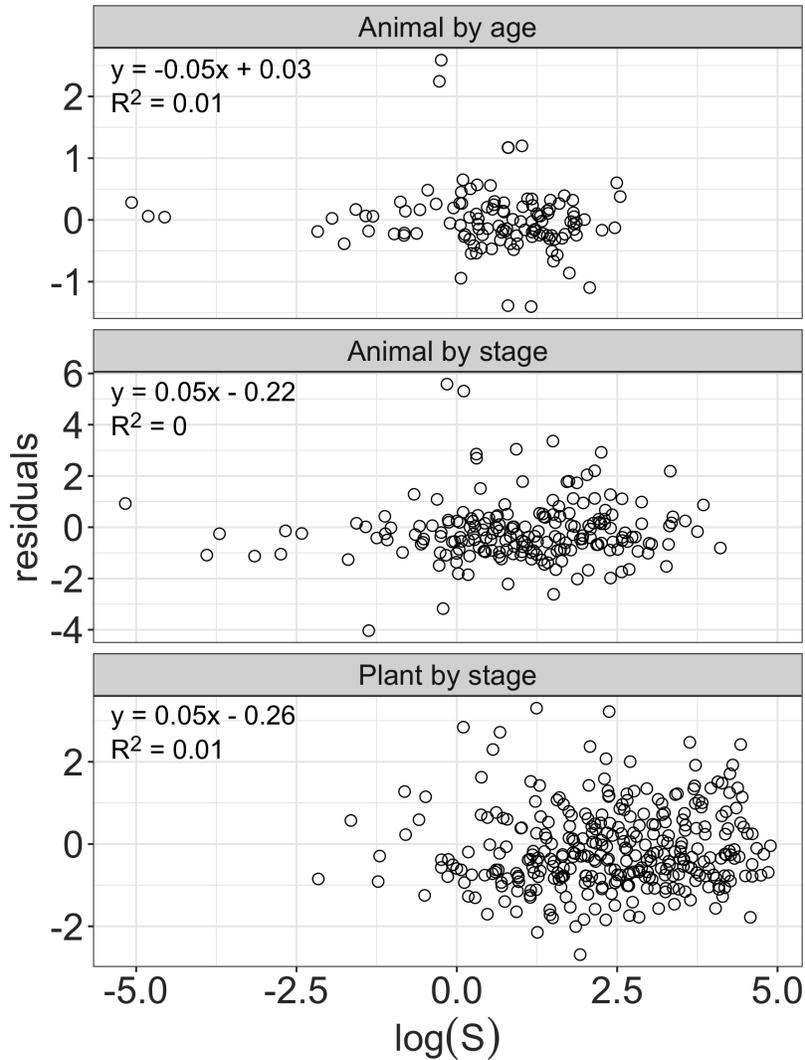


Figure C.6: Residuals of the PGLS model of $\log(\tau)$ and $\log T_c$ versus reproductive dispersion in a log scale ($\log(S)$). On the top left of each panel, we also present the fitted model and its coefficient of determination (R^2) based on ordinary least squares regression (OLS). P-value in each panel is larger than 0.1. The 95% confidence interval for the regression slope is $[-0.09, -0.01]$ for the upper panel, $[-0.01, 0.11]$ for the middle panel and $[0.01, 0.09]$ for the bottom panel.

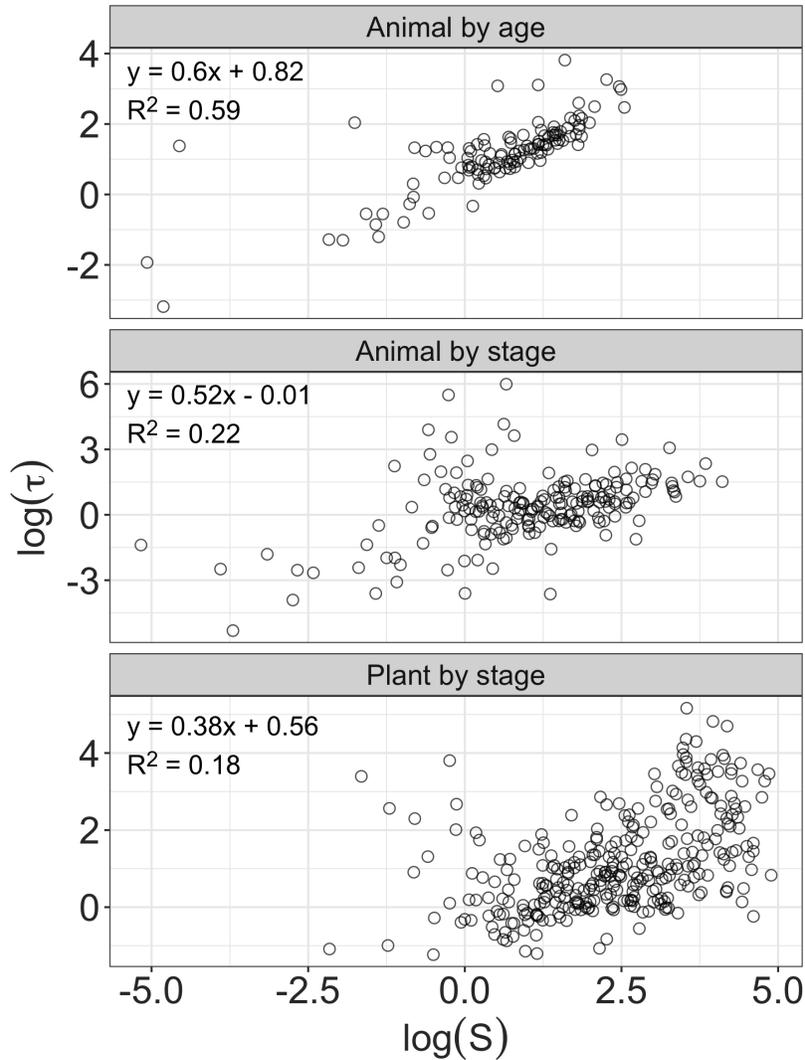


Figure C.7: Damping time (τ) versus reproductive dispersion (S) on a log-log scale. Time unit is years. Upper panel, age-structured animal data from COMADRE and GO; middle panel, stage-structured animal data from COMADRE; bottom panel, stage-structured plant data from COMPADRE. Each panel displays the fitted model and its coefficient of determination (R^2) based on PGLS regression. P-value in each panel is less than 0.001. The 95% confidence interval for the regression slope is [0.56, 0.65] for the upper panel, [0.45, 0.59] for the middle panel and [0.33, 0.42] for the bottom panel. Pagel's λ is 0.00, 0.14 and 0.52 for the upper, middle and bottom panels respectively.

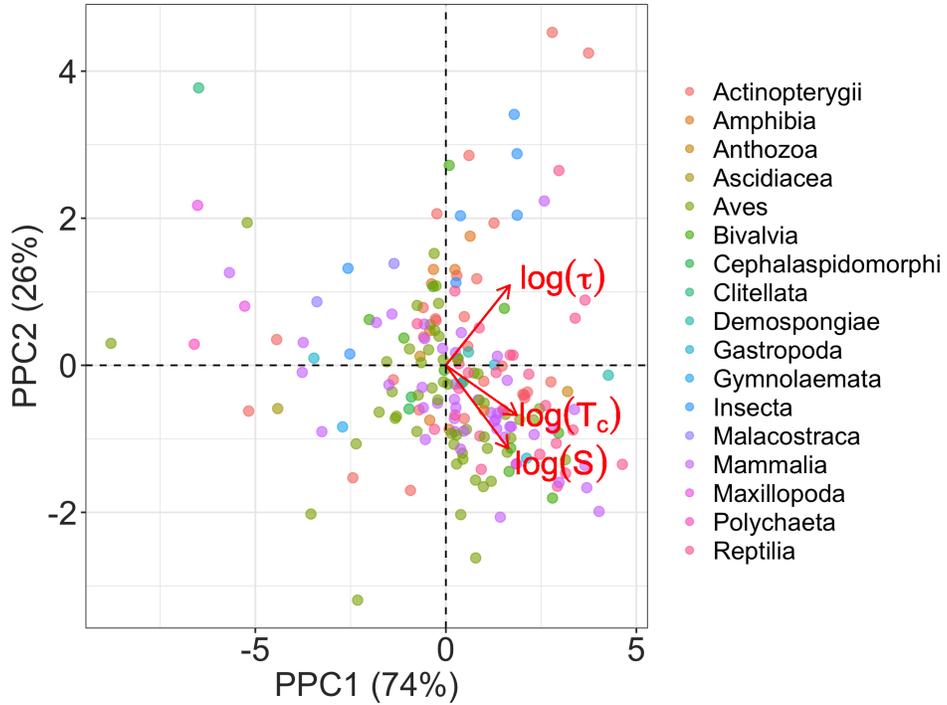


Figure C.8: Phylogenetic principal component analysis (PPCA) of reproductive dispersion (S), generation time (T_c) and damping time (τ) on a log scale for stage-structured animals. Arrow length indicates the loading of each life-history trait onto PCA axes. Points represent the position of species along the phylogenetically corrected principal component (PPC)1 and 2 and are colored by Class. Numbers in parentheses on both axes represent the proportion of variance explained by the corresponding PPC.

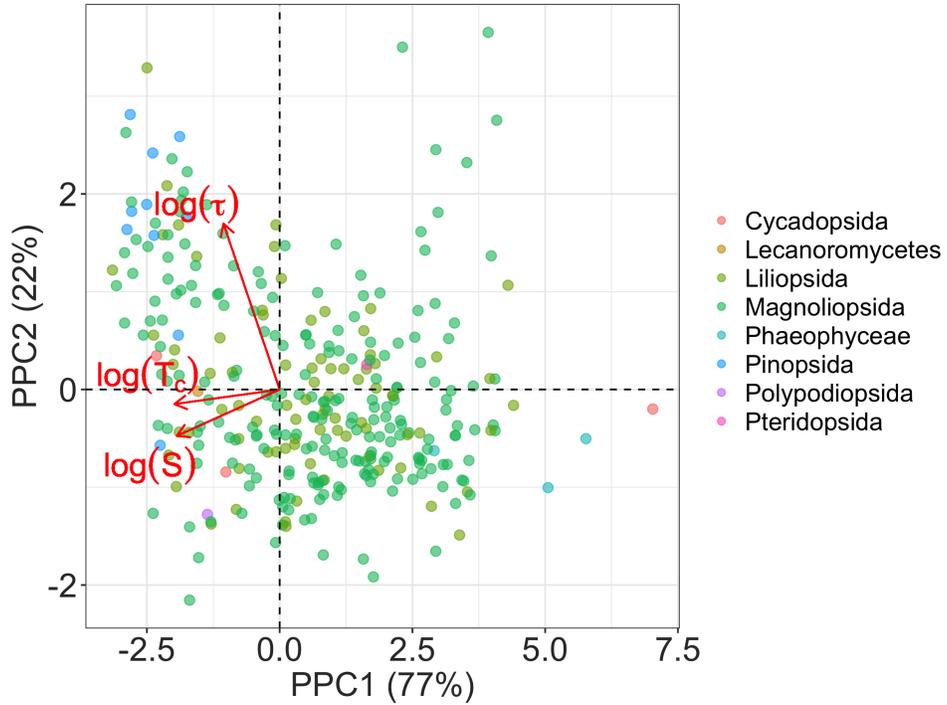


Figure C.9: Phylogenetic principal component analysis (PPCA) of reproductive dispersion (S), generation time (T_c) and damping time (τ) on a log scale for stage-structured plants. Arrow length indicates the loading of each life-history trait onto PCA axes. Points represent the position of species along the phylogenetically corrected principal component (PPC)1 and 2 and are colored by Class. Numbers in parentheses on both axes represent the proportion of variance explained by the corresponding PPC.

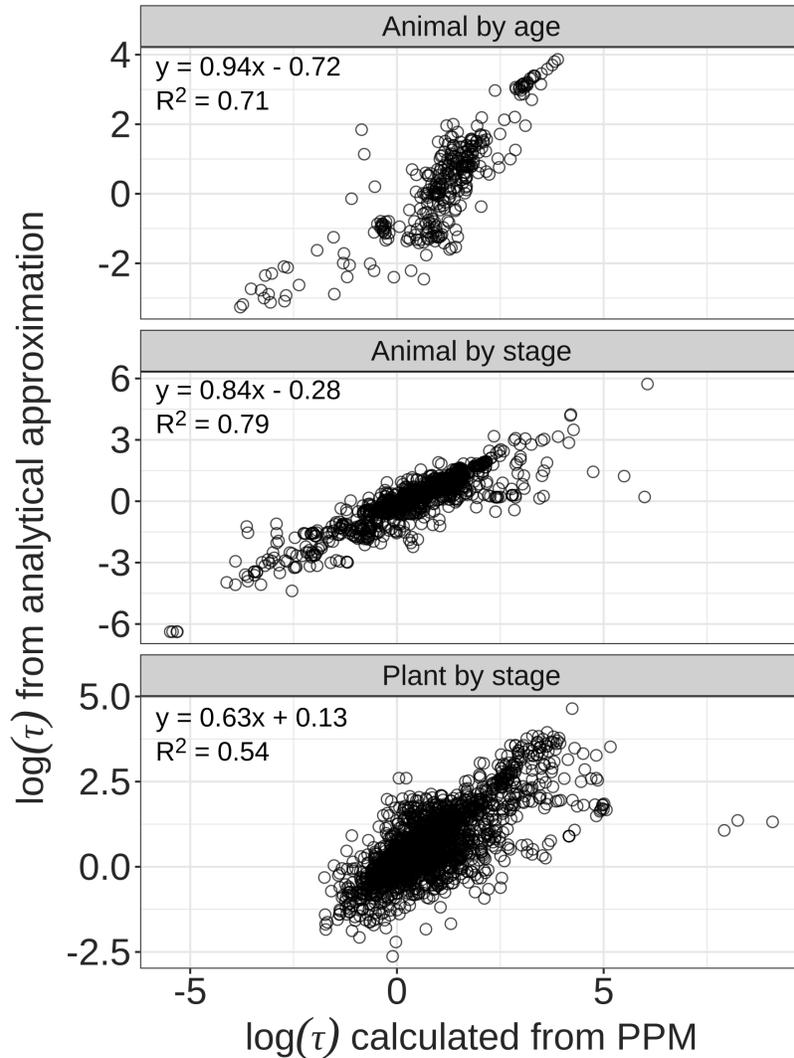


Figure C.10: Comparison between $\log(\tau)$ calculated from population projection matrix (PPM) and $\log(\tau)$ from analytical approximation. Noted that here we include results from 3865 matrices (689 different species). The unit of time is years. Specifically, there are age-structured animal data from COMADRE and GO (*upper*), stage-structured animal data from COMADRE (*middle*), and stage-structured plant data from COMPADRE (*bottom*). On the top left of each panel, we also present the fitted model and its coefficient of determination (R^2) based on ordinary least squares regression (OLS). The 95% confidence interval for the regression slope of each panel is: $[0.91, 0.97]$ for the upper panel, $[0.83, 0.85]$ for the middle panel and $[0.62, 0.64]$ for the bottom panel.